

SCUOLA DI METODOLOGIA DELLA RICERCA CLINICA

2024 - 10^a EDIZIONE

MODULI SPECIALISTICI - S3



GIOVEDÌ 9 MAGGIO 2024

NEGRAR DI VALPOLICELLA (VR)

Centro Formazione IRCCS "Sacro Cuore - Don Calabria"

Generalità e requisiti (G.L. Pappagallo)



Indirect comparisons of competing interventions

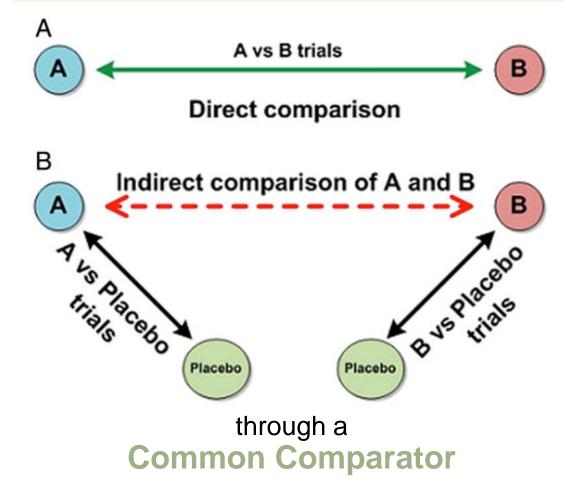
AM Glenny, ^{1*} DG Altman, ² F Song, ³ C Sakarovitch, ² JJ Deeks, ² R D'Amico, ² M Bradburn ² and AJ Eastwood ⁴



Health Technology Assessment 2005; Vol. 9: No. 26

When conducting systematic reviews to evaluate the effectiveness of interventions, direct evidence from good-quality RCTs should be used wherever possible. If little or no such evidence exists, it may be necessary to look for indirect comparisons from RCTs. The reviewer needs, however, to be aware that the results may be susceptible to bias.





Indirect comparisons of competing interventions

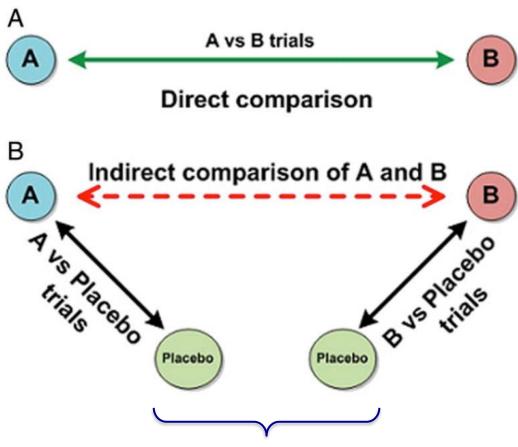
AM Glenny, ^{1*} DG Altman, ² F Song, ³ C Sakarovitch, ² JJ Deeks, ² R D'Amico, ² M Bradburn ² and AJ Eastwood ⁴



Health Technology Assessment 2005; Vol. 9: No. 26

When conducting systematic reviews to evaluate the effectiveness of interventions, direct evidence from good-quality RCTs should be used wherever possible. If little or no such evidence exists, it may be necessary to look for indirect comparisons from RCTs. The reviewer needs, however, to be aware that the results may be susceptible to bias.





Similarity Assumption

trials must be comparable on effect modifiers to obtain an unbiased pooled estimate.

Indirect comparisons of competing interventions

AM Glenny, ^{1*} DG Altman, ² F Song, ³ C Sakarovitch, ² JJ Deeks, ² R D'Amico, ² M Bradburn ² and AJ Eastwood ⁴



Health Technology Assessment 2005; Vol. 9: No. 26

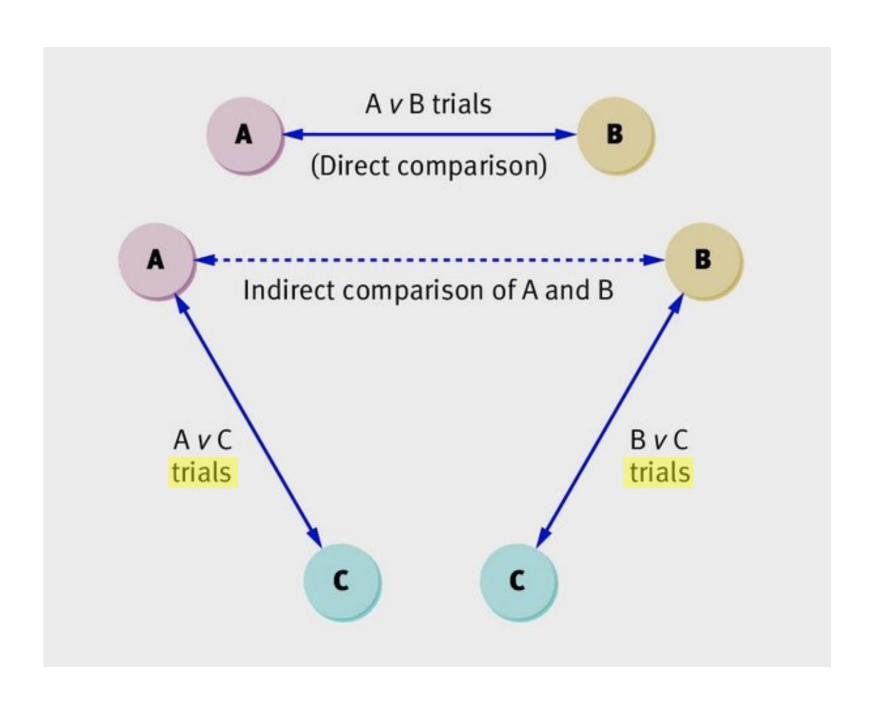
When conducting systematic reviews to evaluate the effectiveness of interventions, direct evidence from good-quality RCTs should be used wherever possible. If little or no such evidence exists, it may be necessary to look for indirect comparisons from RCTs. The reviewer needs, however, to be aware that the results may be susceptible to bias.



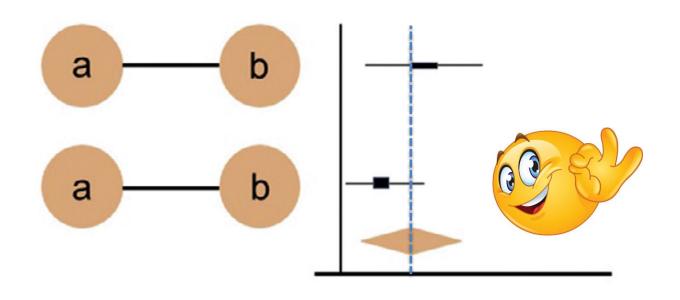
WHAT FACTORS DETERMINE SIMILARITY?

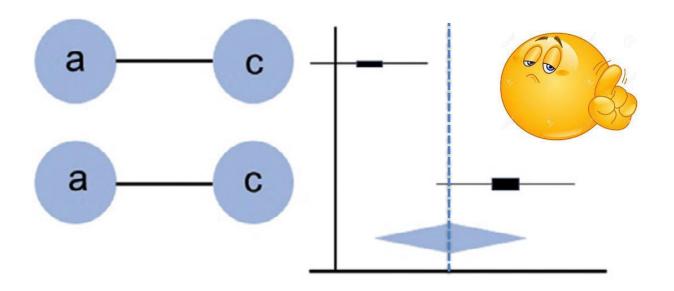
- Included trials should be "comparable" in terms of key factors that could affect the outcome of treatment
- If differences in patient or study characteristics would not be expected to influence treatment effect, the assumption of similarity is not violated
- There are no statistical methods to test for similarity
- Must use clinical knowledge and best judgement to assess appropriate comparability

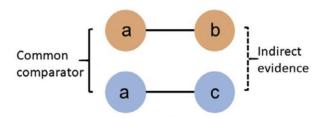




Quando le evidenze dirette sono costituite da più trials...

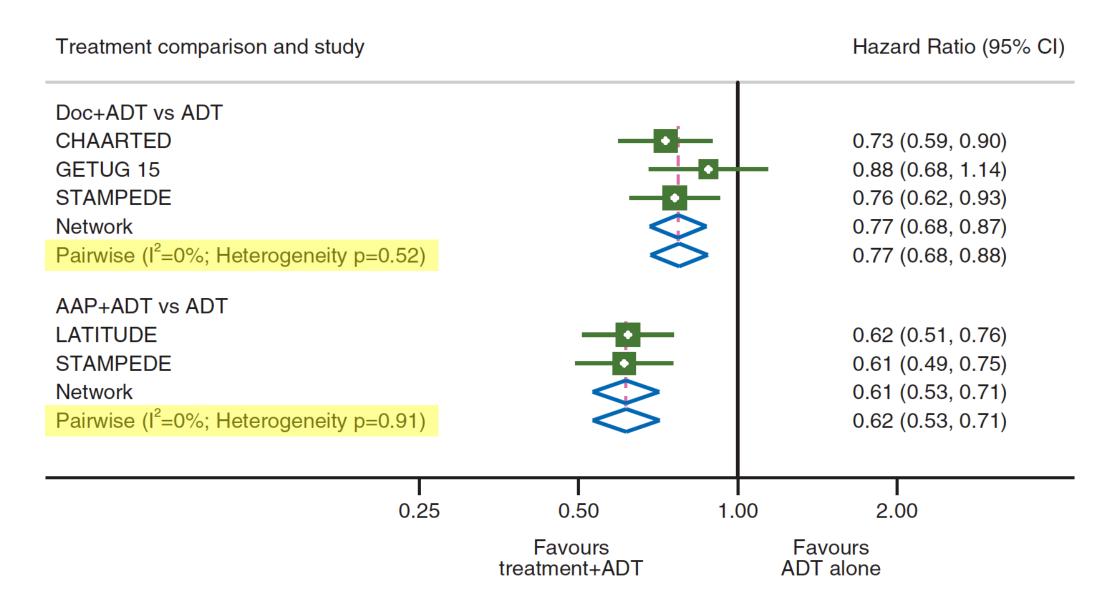






Homogeneity Assumption

there must be no relevant heterogeneity between trial results in pairwise comparisons





Commonly applied methods

Bucher

- IPD not required
- treatment effects calculated for each trial separately
- within study randomization preserved

Population-adjusted indirect comparison (MAIC)

- IPD required for at least 1 trial
- to match the IPD to the AgD of the other trial

Network Meta-Analysis (NMA)

 comparing interventions simultaneously in a single analysis by combining both direct and indirect evidence across a network of studies.



Commonly applied methods

Bucher

- IPD not required
- treatment effects calculated for each trial separately
- within study randomization preserved

Population-adjusted indirect comparison (MAIC)

- IPD required for at least 1 trial
- to match the IPD to the AgD of the other trial

Network Meta-Analysis (NMA)

- comparing interventions simultaneously in a single analysis by combining both direct and indirect evidence across a network of studies.



Commonly applied methods

Bucher

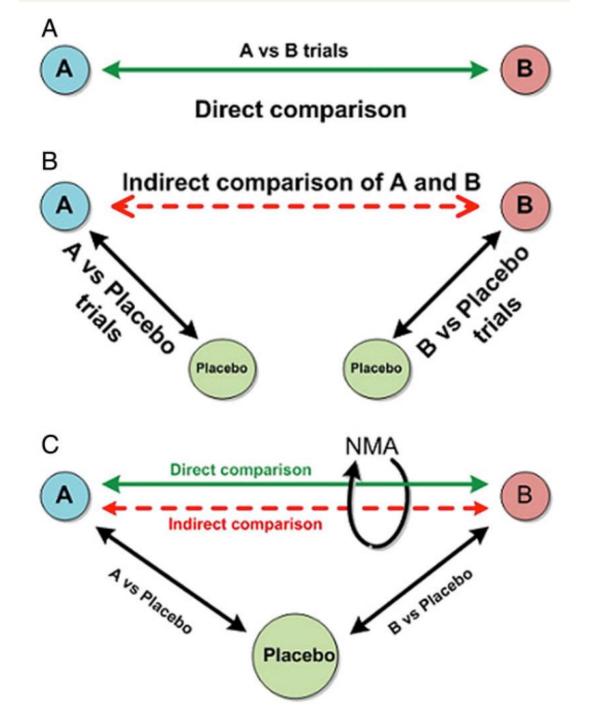
- IPD not required
- treatment effects calculated for each trial separately
- within study randomization preserved

Population-adjusted indirect comparison (MAIC)

- IPD required for at least 1 trial
- to match the IPD to the AgD of the other trial

Network Meta-Analysis (NMA)

 comparing interventions simultaneously in a single analysis by combining both direct and indirect evidence across a network of studies.



Indirect comparisons of competing interventions

AM Glenny, ^{1*} DG Altman, ² F Song, ³ C Sakarovitch, ² JJ Deeks, ² R D'Amico, ² M Bradburn ² and AJ Eastwood ⁴



Health Technology Assessment 2005; Vol. 9: No. 26

When conducting systematic reviews to evaluate the effectiveness of interventions, direct evidence from good-quality RCTs should be used wherever possible. If little or no such evidence exists, it may be necessary to look for indirect comparisons from RCTs. The reviewer needs, however, to be aware that the results may be susceptible to bias.

Consistency Assumption

there must be no relevant discrepancy between direct and indirect evidence





SCUOLA DI METODOLOGIA DELLA RICERCA CLINICA

2024 - 10^a EDIZIONE

MODULI SPECIALISTICI - S3



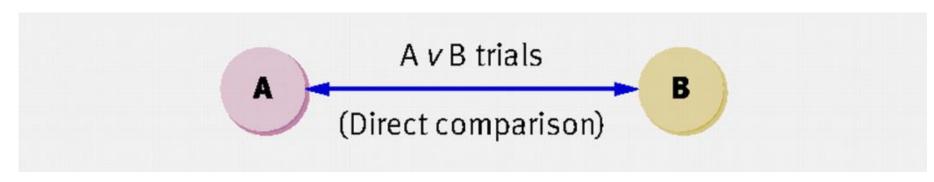
GIOVEDÌ 9 MAGGIO 2024

NEGRAR DI VALPOLICELLA (VR)

Centro Formazione IRCCS "Sacro Cuore - Don Calabria"

Indirect Treatment Comparison (Bucher)

(M. Cinquini)





The best?

No head-to-head comparison







www.elsevier.com/locate/critrevonc

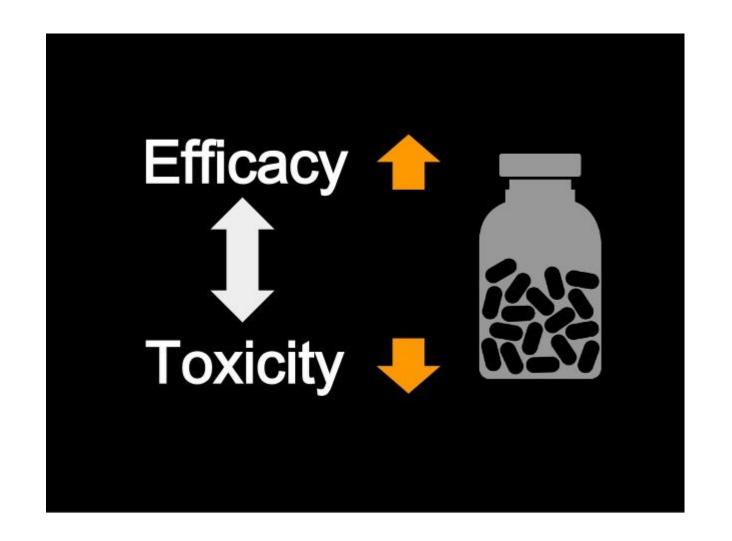
Is there evidence for different effects among EGFR-TKIs? Systematic review and meta-analysis of EGFR tyrosine kinase inhibitors (TKIs) *versus* chemotherapy as first-line treatment for patients harboring EGFR mutations

Eva Regina Haspinger^a, Francesco Agustoni^a, Valter Torri^b, Francesco Gelsomino^a, Marco Platania^a, Nicoletta Zilembo^a, Rosaria Gallucci^a, Marina Chiara Garassino^{a,*}, Michela Cinquini^b

Accepted 11 November 2014

^a Medical Oncology, Fondazione IRCCS Istituto Nazionale dei Tumori, Milan, Italy

^b Fondazione IRCCS Istituto di Ricerche Farmacologiche Mario Negri, Milan, Italy



Population:

- ✓ previously untreated
- ✓ any age and race
- ✓ histologically proven NSCLC harbouring activating EGFR-mutation

Intervention:

✓ EGFR-TKIs (Erlotinib, Gefitinib, Afatinib)

Comparison:

✓ Platinum-based chemotherapy

Outcomes:

- ✓ PFS (whenever possible independently reviewed data)
- ✓ PFS in exon 19 deletion
- ✓ PFS in L858R mutation
- ✓ OS
- ✓ ORR (complete and/or partial and/or stable)
- ✓ Treatment related toxic events

Search strategy

PubMed, Cancer-Lit, Embase-databases and Cochrane-Library were searched for RCTs up to June 2014 with no language or publication status restrictions. Search terms were "TKI" [Substance Name] and "Carcinoma, NSCLC" [Substance Name]. The proceedings of the 2008–2014 conferences of the American Society of Clinical Oncology (ASCO), European Society of Medical Oncology (ESMO) and International Association for the Study of Lung Cancer (IASLC), World Conference of Lung Cancer were also searched for relevant abstracts. Any unpublished RCTs were considered for inclusion.

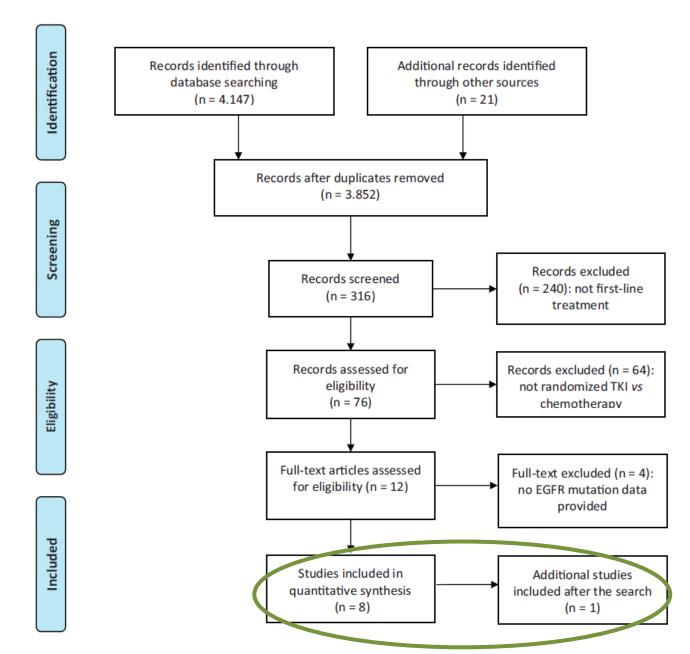


Fig. 1. Flow diagram for the selection of studies included in this meta-analysis.

From: Moher D, Liberati A, Telzlaff J, Altman DG, The PRISMA Group (2009). Preferred reporting items for systematic reviews and meta-analyses: the Prisma statement. PLoS Med 6(6): e1000097. doi:10.1371/journal.pmed1000097.

Indirect Comparisons

Basic assumptions underlying indirect comparisons include:

- homogeneity assumption for standard meta-analysis,
- similarity assumption for adjusted indirect comparison and
- consistency assumption for the combination of direct and indirect evidence. It is essential to fully understand and appreciate these basic assumptions in order to use adjusted indirect and mixed treatment comparisons appropriately.

HOMOGENEITY ASSUMPTION

- When multiple trials are available for a given comparison, the results from multiple trials can be pooled in meta-analyses before an adjusted indirect comparison is conducted.
- For a meta-analysis to be valid, it is commonly established that results from different trials should be sufficiently homogeneous from a clinical and statistical perspective.
- This is usually demonstrated by a 2-tailed p value for homogeneity at Pearson chi-squared test or Cochran Q test > 0.10 and a I² (inconsistency) < 50%.
- When homogeneity is unlikely (e.g. I²>50%) than heterogeneity and inconsistency are likely.

Data synthesis:

- ✓ HR for OS and PFS
- ✓ RR for the Others

OS

Panel B

				Chemotherapy		Hazard Ratio	Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
1.2.1 Gefitinib vs che	emotherapy						
FIRST-SIGNAL	0.0392	0.3755	26	16	6.4%	1.04 [0.50, 2.17]	
IPASS	0	0.143	132	129	44.3%	1.00 [0.76, 1.32]	*
NEJSG002	-0.12	0.171	114	110	31.0%	0.89 [0.63, 1.24]	-
WJT0G3405	0.17	0.223			18.2%		
Subtotal (95% CI)			358	341	100.0%	1.00 [0.83, 1.20]	•
Heterogeneity: Tau2=	= 0.00; Chi ² = 1.08, df	f = 3 (P =	$0.78); I^2 = 0\%$				
Test for overall effect	Z= 0.04 (P = 0.97)						
1.2.2 Erlotinib vs che	emotherapy						
EURTAC	0.039	0.24	86	87	39.5%	1.04 [0.65, 1.66]	
OPTIMAL	0.0677	0.219	82	72	47.4%	1.07 [0.70, 1.64]	*
TORCH	0.457	0.416	19	20	13.1%	1.58 [0.70, 3.57]	+-
Subtotal (95% CI)			187	179	100.0%	1.11 [0.83, 1.50]	*
Heterogeneity: Tau2 =	= 0.00; Chi2 = 0.82, df	f = 2 (P =	0.66); $I^2 = 0\%$				
Test for overall effect	Z = 0.71 (P = 0.48)						
1.2.3 Afatinib							
LUX-LUNG3	0.11	0.22	230	115	100.0%	1.12 [0.73, 1.72]	
Subtotal (95% CI)			230		100.0%	1.12 [0.73, 1.72]	
Heterogeneity: Not ap	pplicable						
Test for overall effect							
							0.01 0.1 1 10 100
Test for subgroup dif	foroncos: Chi² - N 51	df = 2 (D = 0.77\ IZ = 00	v.			Favours TKI-inhibitors Favours Chemotherapy

Test for subgroup differences: $Chi^2 = 0.51$, df = 2 (P = 0.77), $I^2 = 0\%$

PFS

Panel A

			TKI-inhibitors	Chemotherapy		Hazard Ratio	Hazard Ratio
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
1.1.1 Gefitinib vs che	emotherapy				3//-3:	30	
FIRST-SIGNAL	-0.62	0.3584	26	16	11.8%	0.54 [0.27, 1.09]	
IPASS	-0.73	0.146	132	129	32.0%	0.48 [0.36, 0.64]	
NEJSG002	-1.2	0.158	114	110	30.2%	0.30 [0.22, 0.41]	*
WJTOG3405	-0.71	0.189	86	86	26.0%	0.49 [0.34, 0.71]	-
Subtotal (95% CI)			358	341	100.0%	0.43 [0.32, 0.56]	◆
Heterogeneity: Tau ² :	= 0.04; Chi2 = 6.48, df	= 3 (P =	0.09); $I^2 = 54\%$				
Test for overall effect	Z= 6.04 (P < 0.0000	11)					
1.1.2 Erlotinib vs che	emotherapy						
EURTAC	-0.99	0.195	86	87	36.2%	0.37 [0.25, 0.54]	-
OPTIMAL	-1.83	0.233	82	72	34.6%	0.16 [0.10, 0.25]	
TORCH	-0.51	0.354	19	20	29.1%	0.60 [0.30, 1.20]	-
Subtotal (95% CI)			187	179	100.0%	0.32 [0.16, 0.65]	
Heterogeneity: Tau ² :	= 0.32; Chi2 = 12.26, c	df = 2 (P	= 0.002); I ² = 849	%			
Test for overall effect	Z = 3.16 (P = 0.002)						
1.1.3 Afatinib vs che	motherapy						
LUX-LUNG3	-0.545	0.152	230	115	50.6%	0.58 [0.43, 0.78]	■
LUX-LUNG6	-1.27	0.17	242	122	49.4%	0.28 [0.20, 0.39]	
Subtotal (95% CI)			472	237	100.0%	0.41 [0.20, 0.82]	•
	= 0.24; Chi² = 10.11, c	df=1 (P	= 0.001); I ² = 909	%			
Test for overall effect	: Z= 2.49 (P = 0.01)						
							-ttt
							0.005 0.1 1 10 200
Foet for eubaroun dif	ferences: Chi² = 0.55	df- 2/	P = 0.76\ 12 = 0%				Favours TKI-inhibitors Favours Chemotherap

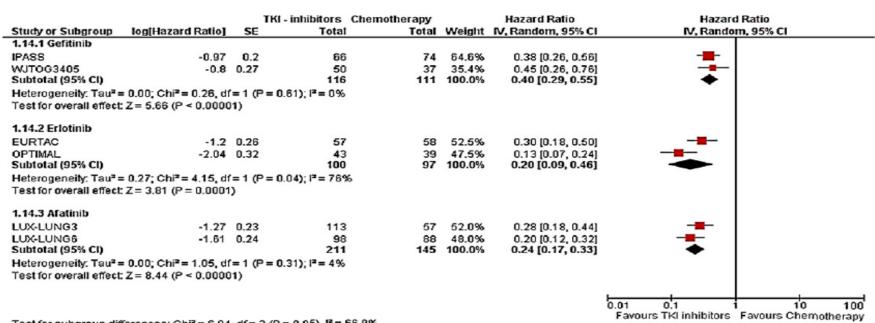
Test for subgroup differences: $Chi^2 = 0.55$, df = 2 (P = 0.76), $I^2 = 0\%$

Exon 21

			TKI-inhibitors	Chemotherapy		Hazard Ratio	Hazard Ratio	
Study or Subgroup	log[Hazard Ratio]	SE	Total	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.15.1 Gefitinib								
IPASS	-0.6	0.23	64	47	61.4%	0.55 [0.35, 0.86]	- -	
WJTOG3405	-0.67	0.29		49	38.6%	0.51 [0.29, 0.90]	 -	
Subtotal (95% CI)			100	96	100.0%	0.53 [0.38, 0.76]	•	
Heterogeneity: Tau2:	= 0.00; Chi ² $= 0.04$, di	f = 1 (F	$P = 0.85$); $I^2 = 0\%$					
Test for overall effect	: Z= 3.48 (P = 0.0005	5)						
1.15.2 Erlotinib								
EURTAC	-0.6	0.32	29	29	50.0%	0.55 [0.29, 1.03]	- = 	
OPTIMAL	-1.35	0.32	39	33	50.0%	0.26 [0.14, 0.49]		
Subtotal (95% CI)			68	62	100.0%	0.38 [0.18, 0.79]	-	
Heterogeneity: Tau ² :	= 0.18; Chi ² = 2.75, dt	f = 1 (F	P = 0.10); $P = 649$	%				
Test for overall effect	: Z = 2.60 (P = 0.009)							
1.15.3 Afatinib								
LUX-LUNG3	-0.31	0.24	91	47	50.7%	0.73 [0.46, 1.17]		
LUX-LUNG6	-1.14	0.26	74	64	49.3%	0.32 [0.19, 0.53]		
Subtotal (95% CI)			165	111	100.0%	0.49 [0.22, 1.10]		
Heterogeneity: Tau2:	= 0.28; Chi ² = 5.50, dt	f = 1 (F	P = 0.02); $P = 829$	%				
Test for overall effect	: Z= 1.73 (P = 0.08)							
								100
Tanking automorphism			0.40 0.70 17				Favours TKI inhibitors Favours chemother	apy

Test for subgroup differences: $Chi^2 = 0.70$, df = 2 (P = 0.70), $I^2 = 0\%$

Exon 19



Test for subgroup differences: Chi² = 6.04, df = 2 (P = 0.05), I^2 = 66.9%

Panel A

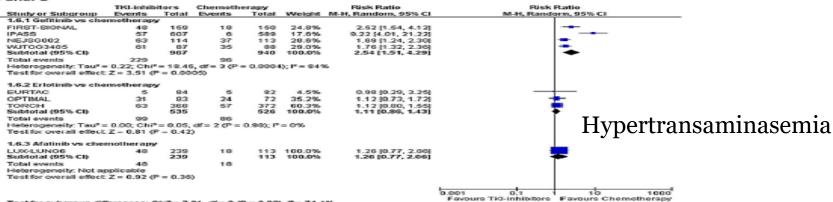
Panel A.									
	TKI-inhib	itors	Chemoth	erapy		Risk Ratio	Risk	Ratio	
Study or Subgroup	Events		Events	Total	Weight	M-H, Random, 95% CI	M-H, Rand	orm, 95% CI	
1.5.1 Gefitinib vs che	motherapy	,							
FIRST-SIGNAL	115	159	20	150	25.1%	5.42 [3.57, 8.25]			
IPASS	402	607	132	589	30.8%	2.96 [2.52, 3.47]		_	
NEJSG002	81	114	25	113		3.21 [2.23, 4.63]			
WJT0G3405	74	87	7	88	17.7%	10.69 [5.22, 21.88]			
Subtotal (95% CI)		967		940	100.0%	4.42 [2.82, 6.92]		•	
Total events	672		184						
Heterogeneity: Tau ² =				= 0.0004	00P = 849	%		l	
Test for overall effect:	Z = 6.47 (F	< 0.DD	001)						
1.5.2 Erlotinib vs cher									Skin reactions
			_						SKIII TEACHOUS
OPTIMAL.	67	84	- 4	82	28.3%	16.35 [6.25, 42.78]			Didiii I Cactions
	61	83	14	72	34.6%	3.78 [2.32, 6.15]			
TORCH Subtotal (95% CI)	252	388 535	135	372 526	37.1%	1.89 [1.82, 2.20] 4.42 [1.57, 12.44]			
		535		320	100.075	4.42[1.37, 12.44]			
Total events	380		153					l	
Heterogeneity: Tau ² =				< 0.0000	(11)(12 = 9)	PS.		l	
Test for overall effect:	2 = 2.81 0	· = 0.00	5)						
1.5.3 afatinib vs chen	notherapy								
LUXGLUNG3	204	229	7	111	40.7%	14.13 [6.89, 28.98]			
LUX-LUNG6	193	239	10	113	59.3%	9.13 [5.03, 16.54]			
Subtotal (95% CI)		468		224	100.0%	10.90 [6.89, 17.24]		•	
Total events	397		17						
Heterogeneity: Tau* =	0.00; Chi*	= 0.85	df = 1 OP =	0.36); 12	= 0%			l	
Test for overall effect:	Z = 10.22	P < 0.0	0001)					l	
							0.001 0.1	10	1000
							Favours TKI-inhibitors		
Test for subgroup diff	erences: C	$hi^* = 8.3$	25, $df = 2$ (1	P = 0.02), $I^a = 75.8$	3 %	Torono Tro-Illinoisora	Cilerinos	a. apj

Panel B

_	anero									
		TKI-inhibi	itors	Chemoth	егару		Risk Ratio	Risk	Ratio	
	Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rand	om, 95% CI	
	1.4.1 Gefitinib vs che	motherapy								
	FIRST-SIGNAL	79	159	45	150	28.0%	1.66 [1.24, 2.21]		-	
	IPASS	283	607	128	589	31.8%	2.15 [1.80, 2.58]			
	NEJ80002	39	114	7	113	13.3%	5.52 [2.58, 11.82]			
	WJTOG3405	47	87	35	88	26.9%	1.36 [0.98, 1.87]		- -	
	Subtotal (95% CI)		967		940	100.0%	2.00 [1.40, 2.85]		-◆	
	Total events	448		215					1	
	Heterogeneity: Tau*=				= 0.0025	; P = 80%			1	
	Test for overall effect:	Z = 3.84 (P)	= 0.00	01)					ı	
	1.4.2 Erlotinib vs che	motherapy	,						l	
	EURTAC	48	84	15	82	35.7%	3.12 [1.91, 5.12]			Diarrhea
	OPTIMAL	21	83	4	72	19,3%	4.55 [1.64, 12.65]			inarmea
	TORCH	152	368	91	372	45.0%	1.69 [1.36, 2.10]		_	Diamina
	Subtotal (95% CI)		535		526	100.0%	2.55 [1.42, 4.56]		-	
	Total events	221		110						
	Heterogeneity: Tau* =				0.02% P	= 75%			1	
	Test for overall effect:	Z = 3.15 0P	0.00	(2)					ı	
	1.4.3 afatinib vs chen	notherapy							l	
	LUX-LUNG3	218	229	17	111	60.0%	6.22 [4.01, 9.64]		_	
	LUX-LUNG5	211	239	12	113	40.0%	8.31 [4.86, 14.22]		=	
	Subtotal (95% CI)		468		224	100.0%	6.98 [4.97, 9.81]			
	Total events	429		29						
	Heterogeneity: Tau ² =	0.00; ChP	= 0.68,	df = 1 QP =	0.41); P	= 0%			1	
	Test for overall effect:	Z = 11.22 (P < 0.0	0001)					ı	
								_	Ι.	_
								0.002 0.1	10	500
		_						Favours TKI-inhibitors	Favours Che	
	To of for outposeum diffi									

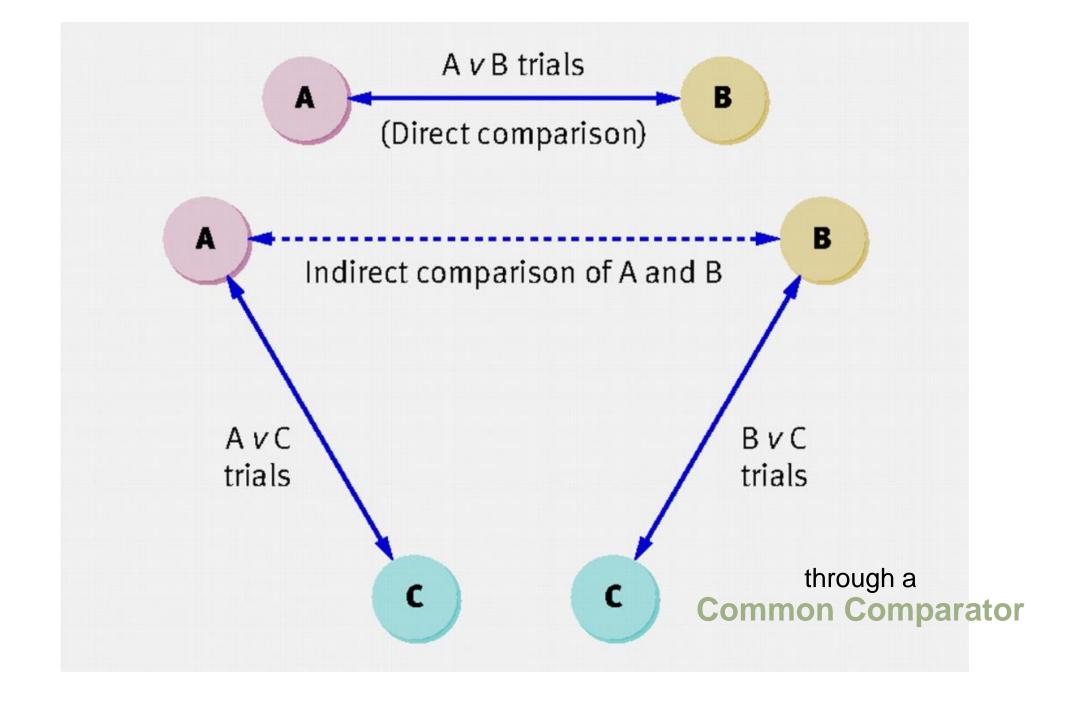
Test for subgroup differences: Chi² = 26.53, df = 2 (P < 0.00001), P = 92.5%

Panel C



SIMILARITY (TRANSITIVITY) ASSUMPTION

- For an adjusted indirect comparison (A vs B) to be valid, a similarity assumption is required in terms of moderators of relative treatment effect.
- That is, patients included should be sufficiently similar in the two sets of control arms (C_1 from the trial comparing A vs C_1 , and C_2 , from the trial comparing B vs C_2).
- This is crucial as only a large theoretical overlap between patients enrolled in C₁ and C₂ enables the relative effect estimated by trials of A versus C₁ to be generalizable to patients in trials of B versus C₁, and the relative effect estimated by trials of B versus C₂ to be generalizable to patients in trials of A versus C₂.



WHAT FACTORS DETERMINE SIMILARITY?

- Included trials should be "comparable" in terms of key factors that could affect the outcome of treatment
- If differences in patient or study characteristics would not be expected to influence treatment effect, the assumption of similarity is not violated
- There are no statistical methods to test for similarity
- Must use clinical knowledge and best judgement to assess appropriate comparability

- Effect modifiers include more than patient characteristics
- Trial and outcome properties are also important
- The PICOS framework can be used to identify these variables

		Would the treatment be expected to work equally in all patients included into the meta-
	Description	Sample Variab analysis?
Р	Patient Population	Demographics, baseline clinical characteristics, disease severity
I	Intervention	Dose, mode of admin, duration
С	Comparator	Active treatment, placebo, concomitant meds
0	Outcomes	Definitions, thresholds, ITT vs. PP, EOS vs. EOT, LOCF vs. NC=F,
S	Setting	Study design, study duration, location/country, method of outcome assessment

- Effect modifiers include more than patient characteristics
- Trial and outcome properties are also important
- The PICOS framework can be used to identify these variables

	Description	Dosing and duration may or may not be important to treatment outcome.					
	Description	Sample Variab be important to treatment outcome.					
Р	Patient Population	Demographics, baseline clinical characteristics, disease severity					
1	Intervention •	Dose, mode of admin, duration					
С	Comparator	Active treatment, placebo, concomitant meds					
0	Outcomes	Definitions, thresholds, ITT vs. PP, EOS vs. EOT, LOCF vs. NC=F,					
S	Setting	Study design, study duration, location/country, method of outcome assessment					

- Effect modifiers include more than patient characteristics
- Trial and outcome properties are also important
- The PICOS framework can be used to identify these variables

	Description	Sample Variables					
Р	Patient Population	Demographics, baseline clinical characteristics, disease severity					
1	Intervention	Dose, mode of admin, duration					
С	Comparator	Active treatment, placebo, concomitant meds					
0	Outcomes	Definitions, thresholds, ITT vs. PP, EOS vs. EOT, LOCF vs. NC=F,					
S	Setting	Study design, study duration, location/country, method of outcome assessment					

In pair-wise meta-analyses the comparator must be the same for each trial. In NMA, the comparators need not be equal, but it must fit within the network.

- Effect modifiers include more than patient characteristics
- Trial and outcome properties are also important
- The PICOS framework can be used to identify these variables

	Description	Sample Variables
Р	Patient Population	Demographics, baseline clinical characteristics, disease severity
1	Intervention	Dose, mode of admin, duration
С	Comparator	Active treatment, placebo, concomitant meds
0	Outcomes •	Definitions, thresholds, ITT vs. PP, EOS vs. EOT, LOCF vs. NC=F,
S	Setting	Study design, study duration, location/country, method of outcome assessment

How outcomes are calculated can influence observed treatment effect.

- Effect modifiers include more than patient characteristics
- Trial and outcome properties are also important
- The PICOS framework can be used to identify these variables

	Description	Sample Variables					
Р	Patient Population	Demographics, baseline clinical characteristics, disease severity					
1	Intervention	Dose, mode of admin, duration					
С	Comparator	Active treatment, placebo, concomitant meds					
0	Outcomes	Definitions, thresholds, ITT vs. PP, EOS vs. EOT, LOCF vs. NC=F,					
S	Setting •	Study design, study duration, location/country, method of outcome assessment					
		Some general study characteristics can be important. Eg, timing of assessments, study locations with different standards of care,					

patient vs. physician-reported outcomes.

Table 1
Characteristics of the 9 clinical trials included in the mata-analysis.

Characteristics of	i the 9 chinical trials incl	ided in the in	ta-anarysis.				
Trial	Primary end-point	TKI	Chemotherapy	Patients (TKI/CT)	EGFR + patients (%)	Asiatic patients (%)	Crossover (%) ^a
IPASS Mok, 2009	Progression-free survival	Gefitinib	Carboplatin + paclitaxel	1.217 (609/608)	21.4	99.8	39.5
WJTOG3405 Mitsudomi, 2010	Progression-free survival	Gefitinib	Cisplatin + paclitaxel	177 (88/89)	100	100	59.3
NEJ002 Maemondo, 2010	Progression-free survival	Gefitinib	Carboplatin + paclitaxel	228 (114/114)	100	100	94.6
First-SIGNAL Han, 2012	Overall survival	Gefitinib	Cisplatin + gemcitabine	309 (159/150)	13.6	100	75.0
TORCH Gridelli, 2012	Overall survival	Erlotinib	Cisplatin + gemcitabine	760 (380/380)	5.1	0	60.9
OPTIMAL Zhou, 2011	Progression-free survival	Erlotinib	Carboplatin + gemcitabine	154 (82/72)	100	100	NA
EURTAC Rosell, 2011	Progression-free survival	Erlotinib	Cisplatin/carboplatin + docetaxel/gemcitabine	173 (86/87)	100	0	76.0
LUX-Lung 3 Sequist, 2012	Progression-free survival	Afatinib	Cisplatin + pemetrexed	345 (230/115)	100	100	75.0
LUX-Lung 6 Wu, 2013	Progression-free survival	Afatinib	Cisplatin + gemcitabine	364 (242/122)	100	100	56.0
^a Patients who	have been treated with	rossover from	chemotherapy to TKI in second	l-line.			

Study		
FIRST-SIGNAL	Cisplatin 75 mg/m ² Gemcitabine 1. ²	ery 3 weeks cles
IPASS	Carboplati mg/mill Paci	veeks up to 6 weeks
NEJG002		
WJTOG3405		to 6 weeks
EURTAC		
OPTIN		v. 4 cycles
TORCI	6/m2	i.v. every 3 weeks up to 6 weeks
LUX-LUNG II	n2 Pemetrexed	i.v. 6 cycles
LUX-LUNG VI	75 mg/m2 Gemcitabine mg/m2 day 1&8	i.v. Up to 6 cycles



So, who's the best?



COMPUTATIONS

 The log relative risk of the adjusted indirect comparison of A and B (lnRR_{A vs B}) can be estimated by:

$$\ln RR_{A \text{ vs } B} = \ln RR_{A \text{ vs } C1} - \ln RR_{B \text{ vs } C2}$$

and its standard error is:

SE (In RR_{A vs B}) =
$$\sqrt{[SE (In RRA vs C1)^2 + SE (In RRB vs C2)^2]}$$

• Similar computations can be envisioned for odds ratio, absolute risk reductions, weighted mean differences, and standardized mean differences.

				Hazard Ratio\Risk ratio	Hazard Ratio\Risk ratio
Study or Subgroup	log[Hazard Ratio\Risk ratio]	SE	Weight	IV, Random, 95% CI	IV, Random, 95% CI
Progression-free survival	0.295	0.385		1.34 [0.63, 2.86]	
PFS-exon 19	0.693	0.447		2.00 [0.83, 4.80]	
PFS-L858R	0.332	0.417		1.39 [0.62, 3.16]	-
Overall survival	-0.104	0.177		0.90 [0.64, 1.27]	+
Objective response rate	-0.036	0.168		0.96 [0.69, 1.34]	+
Diarrhea	-0.223	0.121		0.80 [0.63, 1.01]	+
Rash	0	0.101		1.00 [0.82, 1.22]	+
Hypertransaminasemia	0.83	0.175		2.29 [1.63, 3.23]	(+)
Treatment discontinuation	-0.019	0.384		0.98 [0.46, 2.08]	
Treatment-related death	1.05	1.295		2.86 [0.23, 36.17]	
	[1	lmage o	of Fig. 5		0.05 0.2 1 5 20 Favours Gefitinib Favours Erlotinib

Panel B

Study or Subgroup	log[Hazard Ratio\Risk Ratio]	SE	Weight	Hazard Ratio Risk Ratio N, Random, 95% CI	Hazard Ratio\Risk Ratio IV, Random, 95% CI
Progression-free survival	0.048	0.387		1.05 [0.49, 2.24]	+
PFS-exon 19	0.511	0.235		1.67 [1.05, 2.64]	+
PFS-L858R	0.078	0.447		1.08 [0.45, 2.60]	+
Overall survival	-0.099	0.167		0.91 [0.65, 1.26]	_+
Objective response rate	-0.097	0.157		0.91 [0.67, 1.23]	
Diarrhea	-1.25	0.187		0.29 [0.20, 0.41]	(+)
Rash	-0.903	0.244		0.41 [0.25, 0.65]	+ //
Hypertransaminasemia	0.701	0.276		2.02 [1.17, 3.46]	4
Treatment discontinuation	0.531	0.273		1.70 [1.00, 2.90]	+
Treatment-related death	0.022	0.136		1.02 [0.78, 1.33]	+
					0.001 0.1 1 10 1000 Favours Gelitinib Favours Afatinib

Panel C

Study or Subgroup	log[Hazard Ratio\Risk Ratio]	SF	Weight	Hazard Ratio\Risk Ratio IV, Random, 95% CI	Hazard Ratio\Risk Ratio IV, Random, 95% CI
Progression-free survival	-0.248		recigin	0.78 [0.29, 2.11]	
PFS-exon 19	-0.182			0.83 [0.35, 2.01]	
PFS-L858R	-0.254	0.558		0.78 [0.26, 2.32]	
Objective response rate	-0.061	0.186		0.94 [0.65, 1.35]	+
Overall survival	0.094	0.204		1.10 [0.74, 1.64]	_ +
Hypertransaminasemia	-0.127	0.285		0.88 [0.50, 1.54]	
Diarrhea	-1.01	0.2		0.36 [0.25, 0.54]	+
Rash	-0.903	0.245		0.41 [0.25, 0.66]	+
Treatment discontinuation	0.55	0.395		1.73 [0.80, 3.76]	4
Treatment-related death	-1.03	1.637		0.36 [0.01, 8.83]	
					0.002 0.1 1 10 500 Favours Erlotinib Favours Afatinib

TAKE HOME MESSAGES

 Adjusted indirect comparison meta-analysis represents a simple yet robust tool to make statistical and clinical inference despite the lack of conclusive evidence from head-to-head randomized clinical trials.

• Despite being not at the uppermost level of the hierarchy of evidence based medicine, it can often provide results equivalent to those of subsequent direct comparisons.



SCUOLA DI METODOLOGIA DELLA RICERCA CLINICA

2024 - 10^a EDIZIONE

MODULI SPECIALISTICI - S3

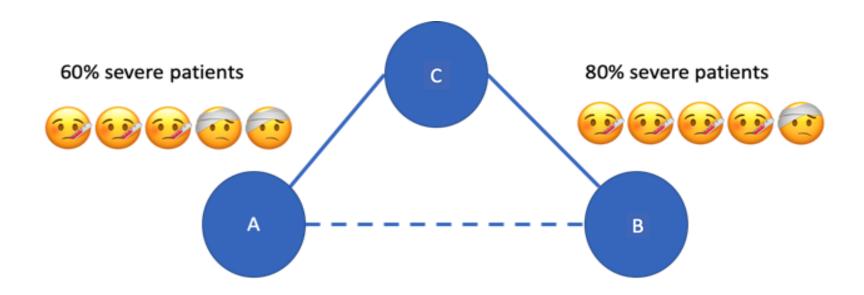


GIOVEDÌ 9 MAGGIO 2024

NEGRAR DI VALPOLICELLA (VR)

Centro Formazione IRCCS "Sacro Cuore - Don Calabria"

Population-Adjusted Indirect Comparison (G.L. Pappagallo)



WHAT FACTORS DETERMINE SIMILARITY?

- Included trials should be "comparable" in terms of key factors that could affect the outcome of treatment
- If differences in patient or study characteristics would not be expected to influence treatment effect, the assumption of similarity is not violated
- There are no statistical methods to test for similarity
- Must use clinical knowledge and best judgement to assess appropriate comparability

Population-adjusted Indirect Comparisons

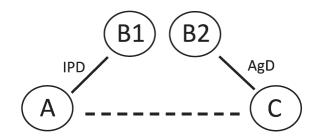
- Population-adjusted Indirect Comparisons use patient-level data from a trial of a given treatment (referred to as the index trial) to derive a comparison of outcomes with competing treatments, based on published information from similarly designed studies, after adjusting for differences in the characteristics of the populations.
- in other words: individual patient data (IPD) in one or more trials are used to adjust for between-trial differences in the distribution of variables that influence outcome
- "anchored" indirect comparison (common comparator arm in each trial) Vs "unanchored" indirect comparison (disconnected treatment network or single-arm studies)
 - an unanchored comparison assumes that all effect modifiers and prognostic factors are accounted for
- unanchored methods for population adjustment are problematic and should not be used when anchored methods can be applied

Population-adjusted Indirect Comparisons

- Population-adjusted Indirect Comparisons use patient-level data from a trial of a given treatment (referred to as the index trial) to derive a comparison of outcomes with competing treatments, based on published information from similarly designed studies, after adjusting for differences in the characteristics of the populations.
- in other words: individual patient data (IPD) in one or more trials are used to adjust for between-trial differences in the distribution of variables that influence outcome
- "anchored" indirect comparison (common comparator arm in each trial) Vs "unanchored" indirect comparison (disconnected treatment network or single-arm studies)
 - an unanchored comparison assumes that all effect modifiers and prognostic factors are accounted for
- unanchored methods for population adjustment are problematic and should not be used when anchored methods can be applied

Matching-adjusted indirect comparison (MAIC)

 needs IPD for at least 1 trial, because the aim is to match the IPD to the AgD of the other trial



- the matching procedure selects a weight for each patient to reach similarity in the summary measures of the baseline characteristics of the IPD and AgD trial and follows the idea of propensity score matching
- the odds between being a patient in trial AB Vs trial CB provides the weights for balancing the populations

Comparative Effectiveness Research in Oncology Methodology: Observational Data

Dawn L. Hershman and Jason D. Wright

J Clin Oncol 30:4215-4222. © 2012 by American Society of Clinical Oncology

Propensity Score Analysis

Propensity score analyses attempt to balance covariates between experimental groups. Using multivariable modeling, the characteristics of a cohort are used to calculate the probability of receiving the intervention. This probability is the propensity score.



MULTIVARIATE ANALYSIS

Independent variables

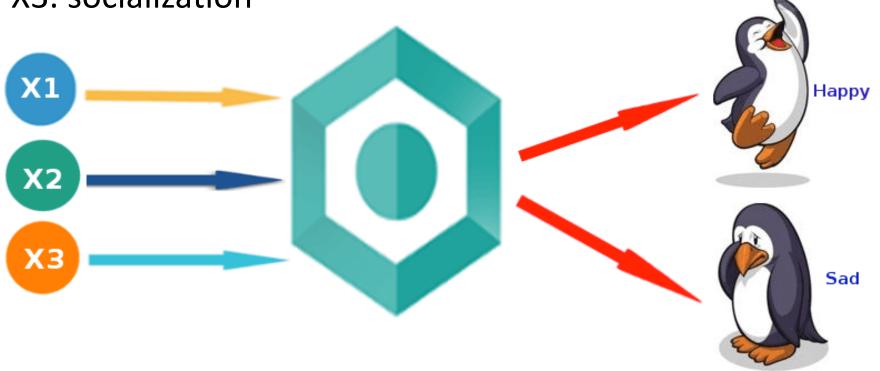
X1. food

X2. water temperature

X3. socialization

Dependent variable

Penguin mood





Multivariate Behavioral Research, 46:399-424, 2011

An Introduction to Propensity Score Methods for Reducing the Effects of Confounding in Observational Studies Peter C. Austin

All measured baseline covariates, all baseline covariates that are associated with treatment assignment, all covariates that affect the outcome (i.e., the potential confounders), and all covariates that affect both treatment assignment and the outcome (i.e., the true confounders).

Propensity Score adjustment

Variabili indipendenti:

Fattori basali tali da influenzare la proposta terapeutica



Variabile dipendente:

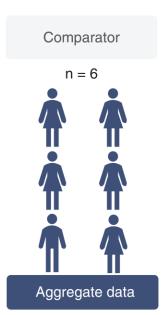
Trattamento assegnato

Matching-adjusted indirect comparison of palbociclib versus ribociclib and abemaciclib in hormone receptor-positive/HER2-negative advanced breast cancer

Hope S Rugo*.¹, Anja Haltner², Lin Zhan³, Anh Tran², Eustratios Bananis³, Becky Hooper², Debanjali Mitra³, & Chris Cameron²

J. Comp. Eff. Res. (2021) 10(6), 457–467



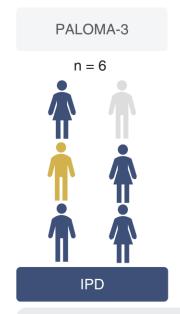


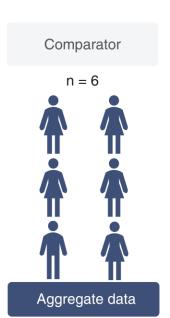
Published trials differ on eligibility criteria and patient characteristics

Matching-adjusted indirect comparison of palbociclib versus ribociclib and abemaciclib in hormone receptor-positive/HER2-negative advanced breast cancer

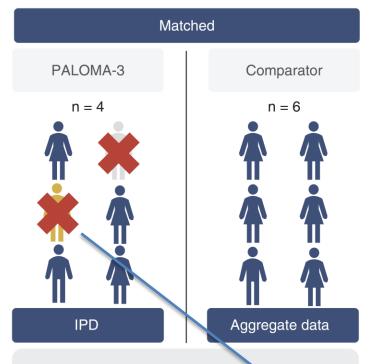
Hope S Rugo*.¹, Anja Haltner², Lin Zhan³, Anh Tran², Eustratios Bananis³, Becky Hooper², Debanjali Mitra³, & Chris Cameron²

J. Comp. Eff. Res. (2021) 10(6), 457-467

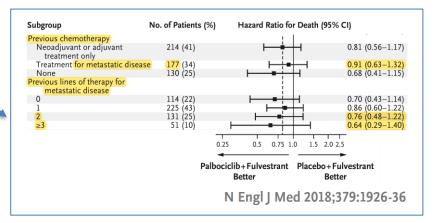




Published trials differ on eligibility criteria and patient characteristics



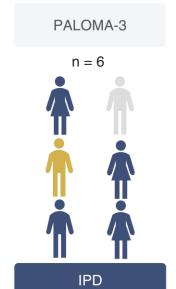
Patients who would not have been eligible for enrolment in the comparator trial are excluded (e.g., patients with prior chemotherapy for MBC, or had ≥2 prior lines of ET for MBC)

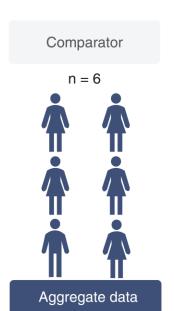


Matching-adjusted indirect comparison of palbociclib versus ribociclib and abemaciclib in hormone receptor-positive/HER2-negative advanced breast cancer

Hope S Rugo*. ¹0, Anja Haltner²0, Lin Zhan³, Anh Tran², Eustratios Bananis³0, Becky Hooper²0, Debanjali Mitra³0 & Chris Cameron²0

J. Comp. Eff. Res. (2021) 10(6), 457-467





Published trials differ on eligibility criteria and patient characteristics

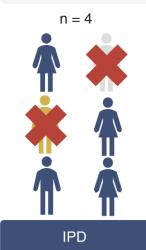
Matched

Comparator

n = 6

Aggregate data

PALOMA-3



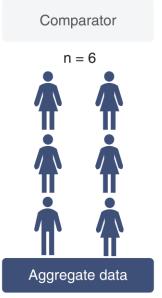
Patients who would not have been eligible for enrolment in the comparator trial are excluded (e.g., patients with prior chemotherapy for MBC, or had ≥2 prior lines of ET for MBC)

Matched and adjusted

PALOMA-3

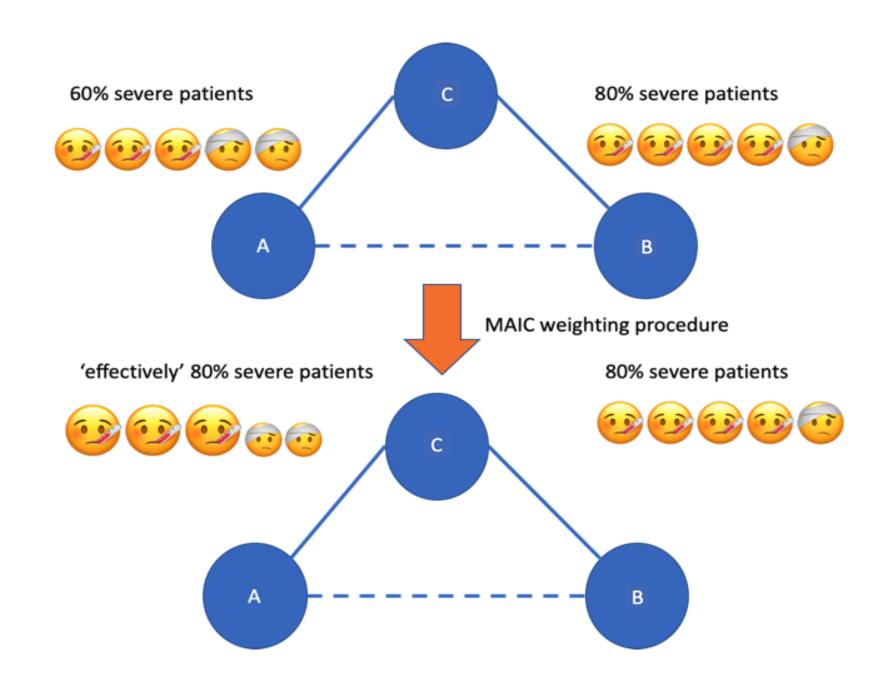
ESS = 3





Patients in PALOMA-3 are weighted to match the averages reported in the comparator trial; ESS reflects practical sample size after adjusting

Adjustment is based on **treatment-effect modifiers** such as prior ET setting and number of lines of therapy for MBC

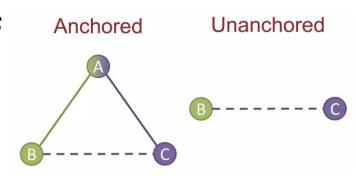




The main limitation relates to the inherent challenge of MAIC in that it is only possible to adjust baseline variables that are mutually reported between trials, and therefore it cannot address the potential unmeasurable differences between the trials.

Population-adjusted Indirect Comparisons

- MAICs and STCs use patient-level data from a trial of a given treatment (referred to as the index trial) to derive a comparison of outcomes with competing treatments, based on published information from similarly designed studies, after adjusting for differences in the characteristics of the populations.
- in other words: individual patient data (IPD) in one or more trials are used to adjust for between-trial differences in the distribution of variables that influence outcome
- "anchored" indirect comparison (common comparator arm in each trial) Vs "unanchored" indirect comparison (disconnected treatment network or single-arm studies)
 - an unanchored MAIC or STC assumes that all effect modifiers and prognostic factors are accounted for
- unanchored methods for population adjustment are problematic and should not be used when anchored methods can be applied





Elranatamab efficacy in MagnetisMM-3 compared with real-world control arms in triple-class refractory multiple myeloma

Luciano J Costa¹, Thomas W LeBlanc², Hans Tesch³, Pieter Sonneveld⁴, Ryan P Kyle⁵, Liliya Sinyavskaya⁵, Patrick Hlavacek⁶, Aster Meche⁶, Jinma Ren⁷, Alex Schepart⁶, Didem Aydin⁶, Guido Nador⁸, Marco daCosta DiBonaventura*, 6

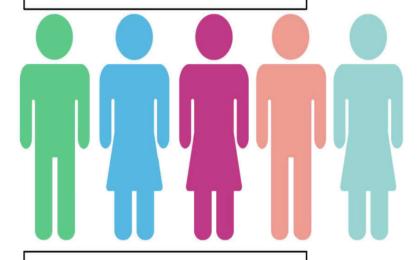
Future Oncol. 2024 Feb 28. doi: 10.2217/fon-2023-0995. Epub ahead of print. PMID: 38415370

Elranatamab efficacy in the single-arm, registrational MagnetisMM-3 trial (NCT04649359) was compared with that of physician's choice of treatment (PCT) for triple-class refractory multiple myeloma. MagnestisMM-3 eligibility criteria were applied to two USA-based oncology electronic health record databases, COTA and Flatiron Health (FH), to identify cohorts for this study (NCT05932290). Applied statistical techniques accounted for cohort imbalances. MagnetisMM-3 (BCMA-naive; n = 123) outcomes were compared with those from COTA (n = 239) and FH (n = 152).

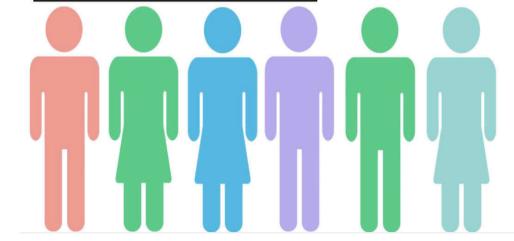


Original, unadjusted population

Clinical trial patients



Real-world patients





Statistical analyses

To summarize our statistical analysis approach, we first controlled for baseline confounding variables by estimating propensity scores (PSs) using age, sex, race, ISS disease stage, ECOG performance status, time from initial MM diagnosis to index date, bone lesions, EMD (COTA database only), high-risk cytogenetics, CCI score, number of LOTs used prior to index, penta-drug refractory status, SCT and levels of aspartate aminotransferase, alanine aminotransferase, hemoglobin, creatinine clearance, calcium, bilirubin and serum albumin. PSs were then used to calculate inverse probability of treatment (IPT) weights, which balanced the distributions of these confounding variables across treatment groups.

The propensity score gives the probability of an individual being exposed (i.e. assigned to the intervention or risk factor) given their baseline characteristics.

The aim of the propensity score in observational research is to control for measured confounders by achieving balance in characteristics between exposed and unexposed groups.

Inverse probability of treatment weighting (IPTW) can be used to adjust for confounding in observational studies.

IPTW uses the propensity score to balance baseline patient characteristics in the exposed and unexposed groups by weighting each individual in the analysis by the inverse probability of receiving his/her actual exposure.



Statistical analyses

To summarize our statistical analysis approach, we first controlled for baseline confounding variables by estimating propensity scores (PSs) using age, sex, race, ISS disease stage, ECOG performance status, time from initial MM diagnosis to index date, bone lesions, EMD (COTA database only), high-risk cytogenetics, CCI score, number of LOTs used prior to index, penta-drug refractory status, SCT and levels of aspartate aminotransferase, alanine aminotransferase, hemoglobin, creatinine clearance, calcium, bilirubin and serum albumin. PSs were then used to calculate inverse probability of treatment (IPT) weights, which balanced the distributions of these confounding variables across treatment groups.

The propensity score gives the probability of an individual being exposed (i.e. assigned to the intervention or risk factor) given their baseline characteristics.

The aim of the propensity score in observational research is to control for measured confounders by achieving balance in characteristics between exposed and unexposed groups.

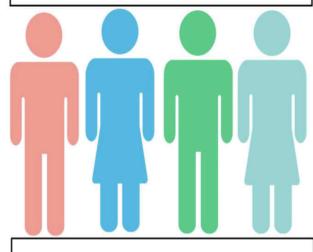
Inverse probability of treatment weighting (IPTW) can be used to adjust for confounding in observational studies.

IPTW uses the propensity score to balance baseline patient characteristics in the exposed and unexposed groups by weighting each individual in the analysis by the inverse probability of receiving his/her actual exposure.

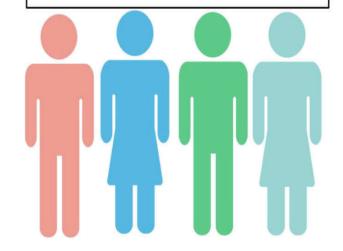


Adjusted population after PS matching

Retained clinical trial patients



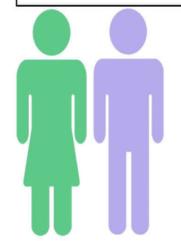
Retained real-world patients



Dropped clinical trial patient



Dropped real-word patients



Patients in the clinical trial get "paired up" with a similar real-world patient. Patients that don't find a match get dropped.



Statistical analyses

To summarize our statistical analysis approach, we first controlled for baseline confounding variables by estimating propensity scores (PSs) using age, sex, race, ISS disease stage, ECOG performance status, time from initial MM diagnosis to index date, bone lesions, EMD (COTA database only), high-risk cytogenetics, CCI score, number of LOTs used prior to index, penta-drug refractory status, SCT and levels of aspartate aminotransferase, alanine aminotransferase, hemoglobin, creatinine clearance, calcium, bilirubin and serum albumin. PSs were then used to calculate inverse probability of treatment (IPT) weights, which balanced the distributions of these confounding variables across treatment groups.

The propensity score gives the probability of an individual being exposed (i.e. assigned to the intervention or risk factor) given their baseline characteristics.

The aim of the propensity score in observational research is to control for measured confounders by achieving balance in characteristics between exposed and unexposed groups.

Inverse probability of treatment weighting (IPTW) can be used to adjust for confounding in observational studies.

IPTW uses the propensity score to balance baseline patient characteristics in the exposed and unexposed groups by weighting each individual in the analysis by the inverse probability of receiving his/her actual exposure.

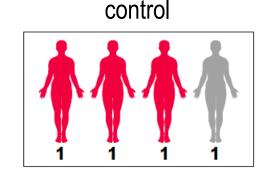
- probability of High-Risk (red figures) in treatment arm: 25%
- inverse probability weight of High-Risk in treatment arm: 1/0.25 = 4
- inverse probability weight of High-Risk in control arm: 1/(1-0.25) = 1.33
- probability of Low-Risk (grey figures) in treatment arm: 75%
- inverse probability weight of Low-Risk in treatment arm: 1/0.75 = 1.33
- inverse probability weight of Low-Risk in control arm: 1/(1-0.75) = 4

After applying the inverse probability weights to create weighted pseudopopulation, the High-Risk characteristic is equally distributed across treatment groups (50% in each group).

Clinical Kidney Journal, 2022, vol. 15, no. 1, 14–20

Original sample

Weighted sample



treatment





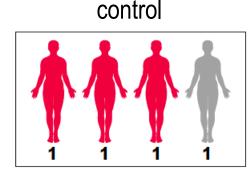
- probability of High-Risk (red figures) in treatment arm: 25%
- inverse probability weight of High-Risk in treatment arm: 1/0.25 = 4
- inverse probability weight of High-Risk in control arm: 1/(1-0.25) = 1.33
- probability of Low-Risk (grey figures) in treatment arm: 75%
- inverse probability weight of Low-Risk in treatment arm: 1/0.75 = 1.33
- inverse probability weight of Low-Risk in control arm: 1/(1-0.75) = 4

After applying the inverse probability weights to create weighted pseudopopulation, the High-Risk characteristic is equally distributed across treatment groups (50% in each group).

Original sample

Weighted

sample



treatment





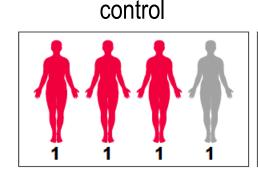
Clinical Kidney Journal, 2022, vol. 15, no. 1, 14–20

- probability of High-Risk (red figures) in treatment arm: 25%
- inverse probability weight of High-Risk in treatment arm: 1/0.25 = 4
- inverse probability weight of High-Risk in control arm: 1/(1-0.25) = 1.33
- probability of Low-Risk (grey figures) in treatment arm: 75%
- inverse probability weight of Low-Risk in treatment arm: 1/0.75 = 1.33
- inverse probability weight of Low-Risk in control arm: 1/(1-0.75) = 4

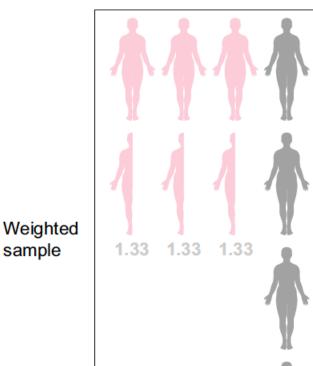
After applying the inverse probability weights to create weighted pseudopopulation, the High-Risk characteristic is equally distributed across treatment groups (50% in each group).

Original sample

sample



treatment





Clinical Kidney Journal, 2022, vol. 15, no. 1, 14–20

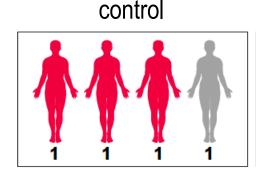
- probability of High-Risk (red figures) in treatment arm: 25%
- inverse probability weight of High-Risk in treatment arm: 1/0.25 = 4
- inverse probability weight of High-Risk in control arm: 1/(1-0.25) = 1.33
- probability of Low-Risk (grey figures) in treatment arm: 75%
- inverse probability weight of Low-Risk in treatment arm: 1/0.75 = 1.33
- inverse probability weight of Low-Risk in control arm: 1/(1-0.75) = 4

After applying the inverse probability weights to create pseudopopulation, the weighted High-Risk characteristic is equally distributed across treatment groups (50% in each group).

Original sample

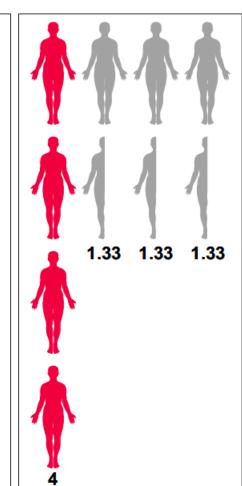
Weighted

sample



treatment



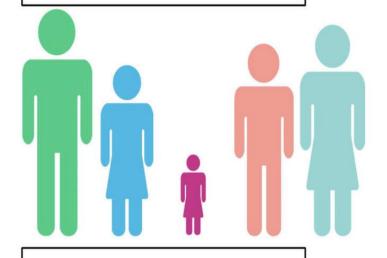


Clinical Kidney Journal, 2022, vol. 15, no. 1, 14–20

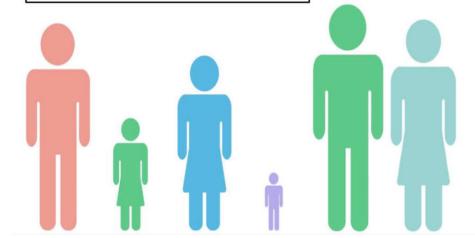


Adjusted population after PS weighting

Clinical trial patients



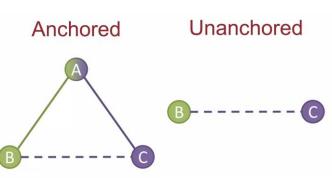
Real-world patients



More similar
patients in each
group get
up-weighted, while
the less similar
patients get
down-weighted. All
patients are
retained.

Population-adjusted Indirect Comparisons

- MAICs and STCs use patient-level data from a trial of a given treatment (referred to as the index trial) to derive a comparison of outcomes with competing treatments, based on published information from similarly designed studies, after adjusting for differences in the characteristics of the populations.
- in other words: individual patient data (IPD) in one or more trials are used to adjust for between-trial differences in the distribution of variables that influence outcome
- "anchored" indirect comparison (common comparator arm in each trial) Vs "unanchored" indirect comparison (disconnected treatment network or single-arm studies)
 - an unanchored MAIC or STC assumes that all effect modifiers and prognostic factors are accounted for
- unanchored methods for population adjustment are problematic and should not be used when anchored methods can be applied





SCUOLA DI METODOLOGIA DELLA RICERCA CLINICA

2024 - 10^a EDIZIONE

MODULI SPECIALISTICI - S3



GIOVEDÌ 9 MAGGIO 2024

NEGRAR DI VALPOLICELLA (VR)

Centro Formazione IRCCS "Sacro Cuore - Don Calabria"

Network Meta-Analysis (NMA)

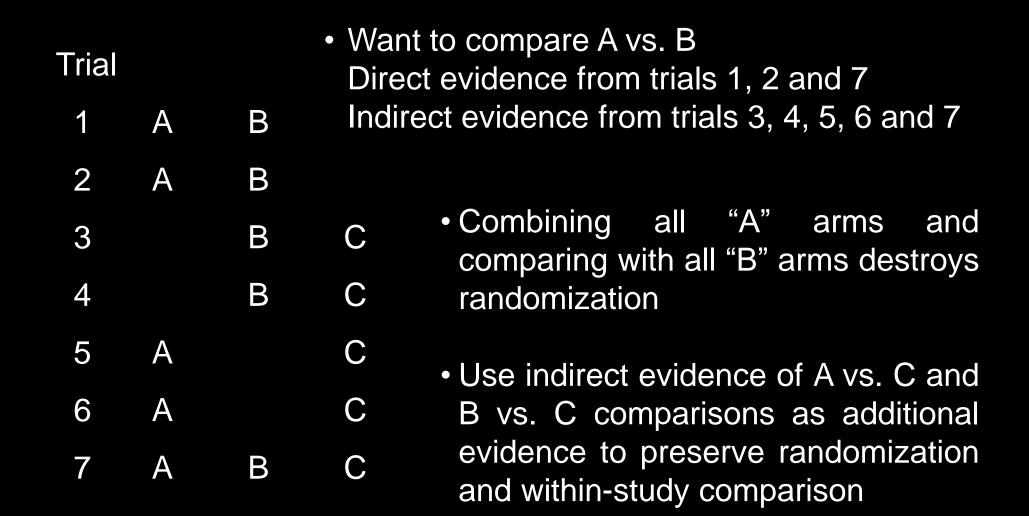
(M. Cinquini)

Network Meta-Analysis

(Multiple Treatments Meta-Analysis, Mixed Treatment Comparisons)

- Combine direct + indirect estimates of multiple treatment effects
- Internally consistent set of estimates that respects randomization
- Estimate effect of each intervention relative to every other whether or not there is direct comparison in studies
- Calculate probability that each treatment is most effective
- Compared to conventional pair-wise meta-analysis:
 - Greater precision in summary estimates
 - Ranking of treatments according to effectiveness

Indirect Comparisons of Multiple Treatments – Network Meta-Analysis



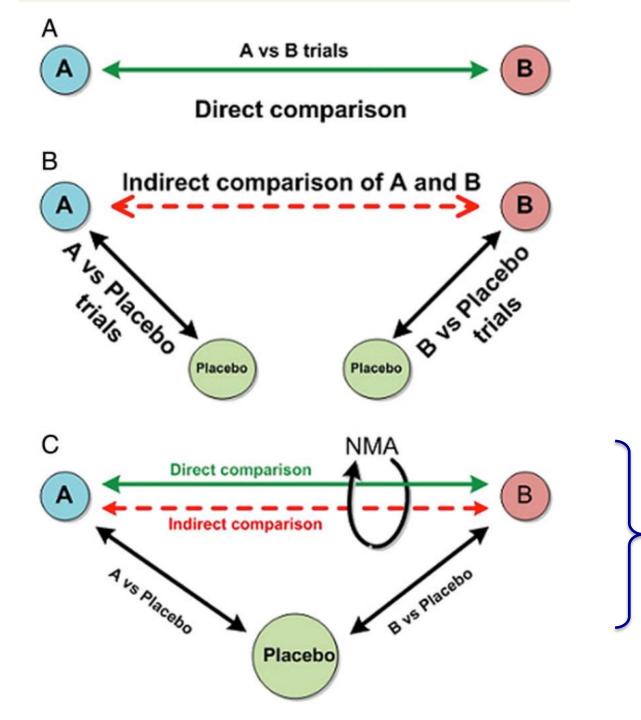
Indirect Comparisons

Basic assumptions underlying indirect comparisons include:

- ✓ homogeneity assumption for standard meta-analysis,
- ✓ **similarity** assumption for adjusted indirect comparison and
- consistency assumption for the combination of direct and indirect evidence. It is essential to fully understand and appreciate these basic assumptions in order to use adjusted indirect and mixed treatment comparisons appropriately.

CONSISTENCY ASSUMPTION

- When both direct and indirect evidence is available, an assumption of evidence consistency is required to quantitatively combine the direct and indirect estimates.
- It is important to investigate possible causes of discrepancy between the direct and indirect evidence, such as the play of chance, invalid indirect comparison, bias in head-to-head comparative trials, and clinically meaningful heterogeneity
- When the direct comparison differs from the adjusted indirect comparison, we should usually give more credibility to evidence from head-to-head comparative trials. However, evidence from direct comparative trials may not always be valid.

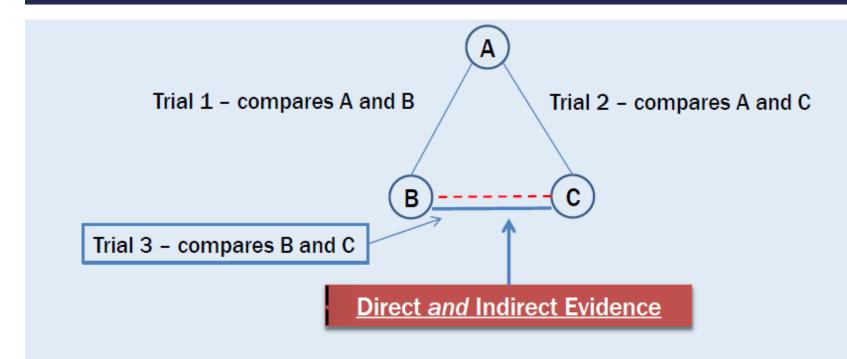


Consistency Assumption

there must be no relevant discrepancy between direct and indirect evidence



THERE ARE 2 TYPES OF TRIAL EVIDENCE



Consistency ⇒ **Direct and indirect evidence agree**

Inconsistency ⇒ Direct and indirect evidence disagree

Differing effect modifiers among the trials can cause inconsistency

METHODS TO TEST FOR INCONSISTENCY

1. Bucher method

- Can be used on triangle structures where three direct estimates are available
- All such "triangles" should be evaluated one by one

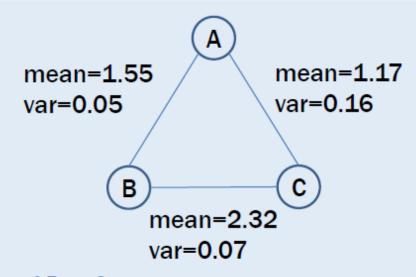
2. Node-splitting

- Direct and indirect studies are separated and a difference in estimates is calculated
- Repeated for all treatment comparisons where inconsistency is possible

3. Inconsistency model

- Could be considered "independence" model because all treatment comparisons are estimated independently
- Treatment effects are not estimated relative to a reference treatment

#1 BUCHER METHOD ILLUSTRATION



Indirect estimate of B vs C:

=
$$1.17$$
 (A vs C) - 1.55 (A vs B) = -0.38 variance = $0.16 + 0.05 = 0.21$

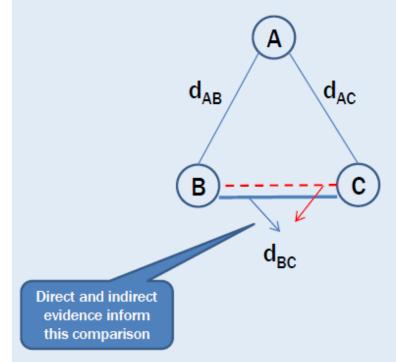
Measure of inconsistency (Z):

If $Z/\sqrt{Var(Z)}$ is rejected (N(0,1)) then the loop is inconsistent

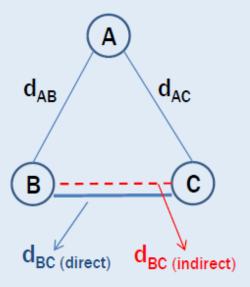
In this case P<.00001, indicating inconsistency

#2 NODE-SPLITTING

Full NMA estimates 3 parameters



Node-splitting estimates separate parameters for direct and indirect evidence

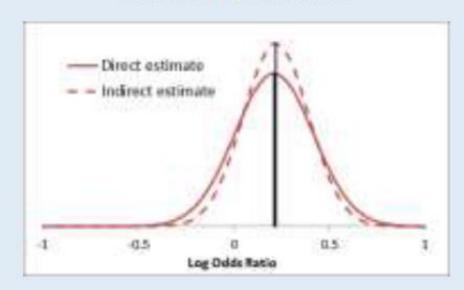


Inconsistency is present if $d_{BC \; (direct)} \neq d_{BC \; (indirect)}$

#2 NODE-SPLITTING

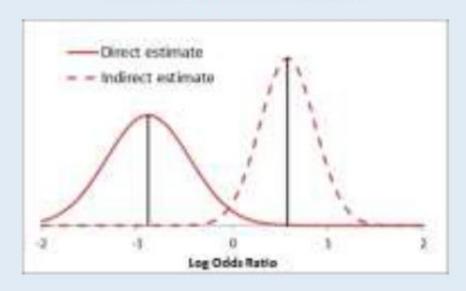
Example of posterior distributions with direct and indirect evidence

Consistent Evidence



Posterior densities overlap indicating absence of inconsistency

Inconsistent Evidence

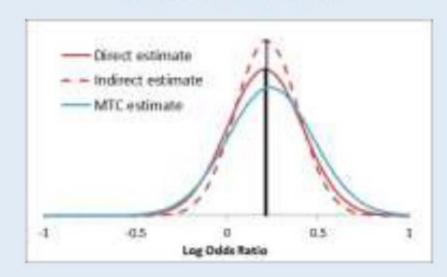


Posterior densities hardly overlap indicating presence of inconsistency

#2 NODE-SPLITTING

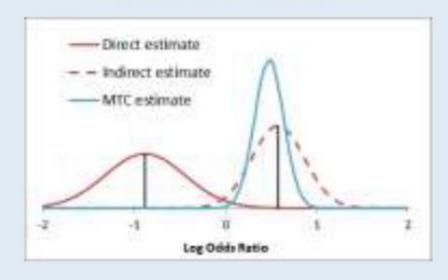
What do we do with this information?

Consistent Evidence



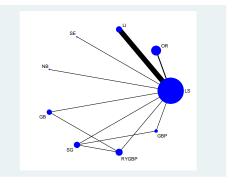
MTC estimate is similar to direct and indirect estimates

Inconsistent Evidence



MTC estimate is closer to indirect estimate, possibly because indirect trials are larger and more precise

Step 1: generating network geometry



Step 3: creating plots and league table of effect size by treatment

Table 1. Inconsistency test between direct and indirect treatment comparisons in mixed treatment comparison

Side	Dire	ct	Indire	ect	Differe	D. 7	
	Coefficient	SE	Coefficient	SE	Coefficient	SE	- p>z
AB	-1.083	0.174	-0.877	0.620	-0.206	0.636	0.746
AC	-1.388	0.247	-1.869	0.493	0.481	0.542	0.375
A D	-1.378	0.265	-0.738	0.413	-0.640	0.479	0.182
ΑE	-3.425	0.940	-3.221	1.005	-0.204	0.937	0.828
BC	-0.894	0.655	-0.312	0.297	-0.581	0.715	0.416
BD	0.099	0.462	-0.241	0.329	0.340	0.567	0.548
BE	-2.152	0.881	-2.615	1.087	0.463	0.896	0.605
CD	0.490	0.492	0.177	0.350	0.313	0.604	0.605
DE	-2.550	1.254	-1.956	0.958	-0.595	1.314	0.651

SE, standard error; A, placebo; B, IV_single; C, IV_double; D, topical; E, combination.

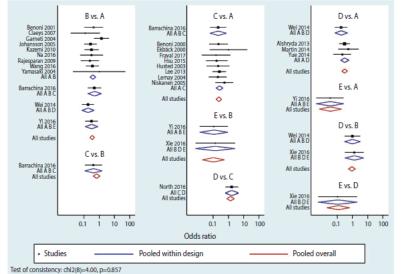
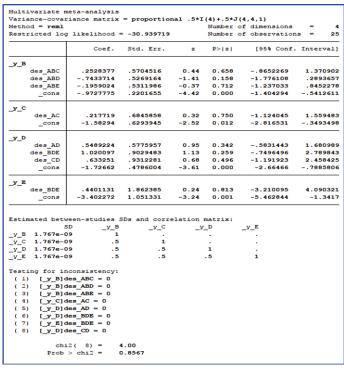


Figure 4. Network forest plot. A, placebo; B, IV_single; C, IV_double; D, topical; E, combination.

Step 2: testing for inconsistency



Step 4: determining relative rankings of treatment

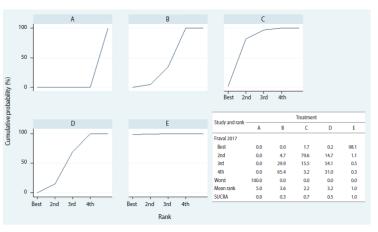


Figure 6. Results of network rank test. A, placebo; B, IV_single; C, IV_double; D, topical; E, combination; SCURA, surface under the cumulative ranking.

Presenting the data

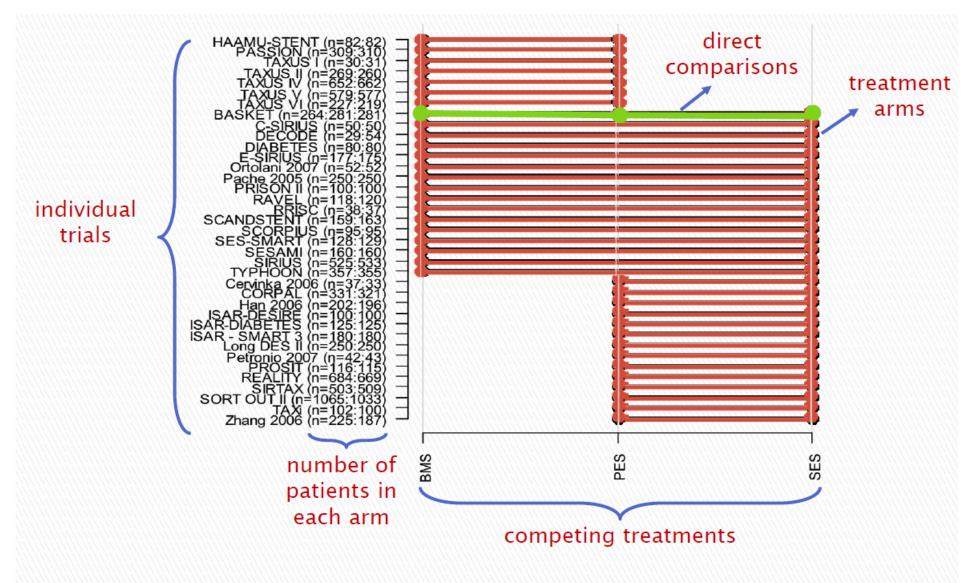
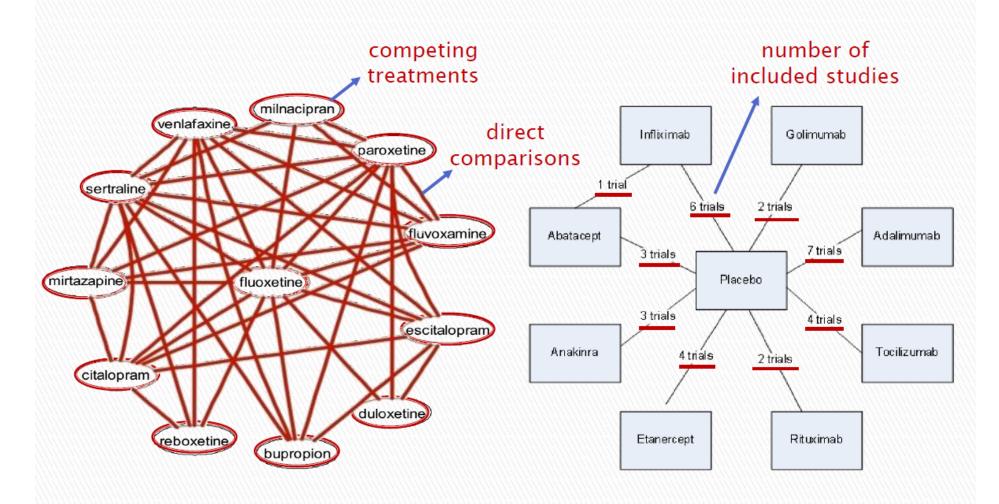


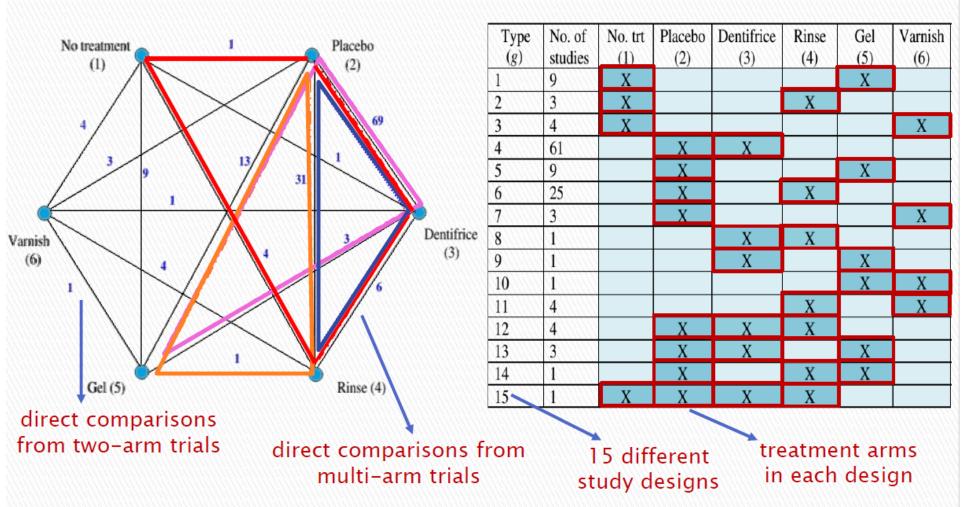
Diagram showing the comparisons involved in the individual studies of the network

[Example in Hoaglin et al. 2011]



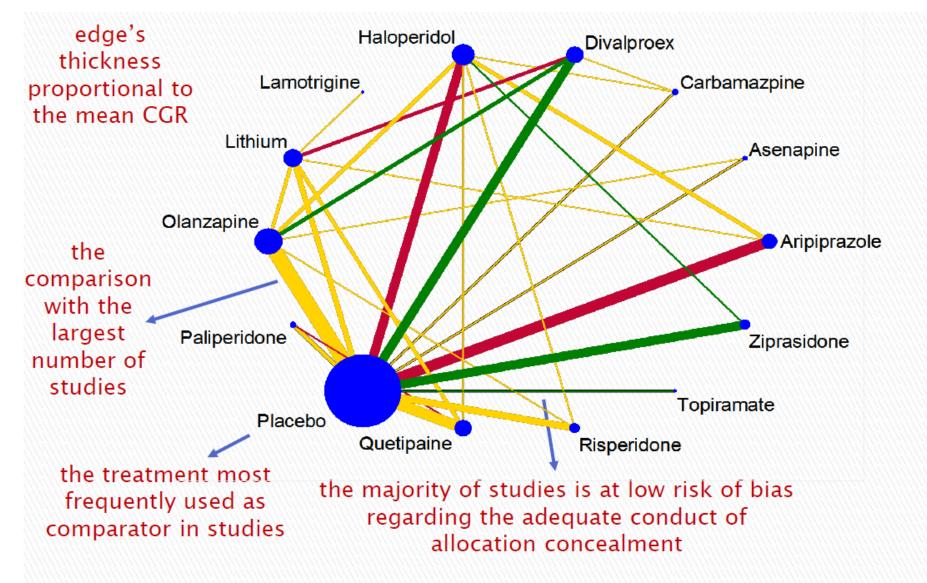
Network graph showing the available direct comparisons in the network

[Examples in Hoaglin et al. 2011 & Jonas et al. 2013]



Network graph showing the presence of multi-arm trials & table showing the network structure; the available study designs in the network

[Examples in Lu et al. 2011]



Network graph with weighted and/or colored nodes and edges

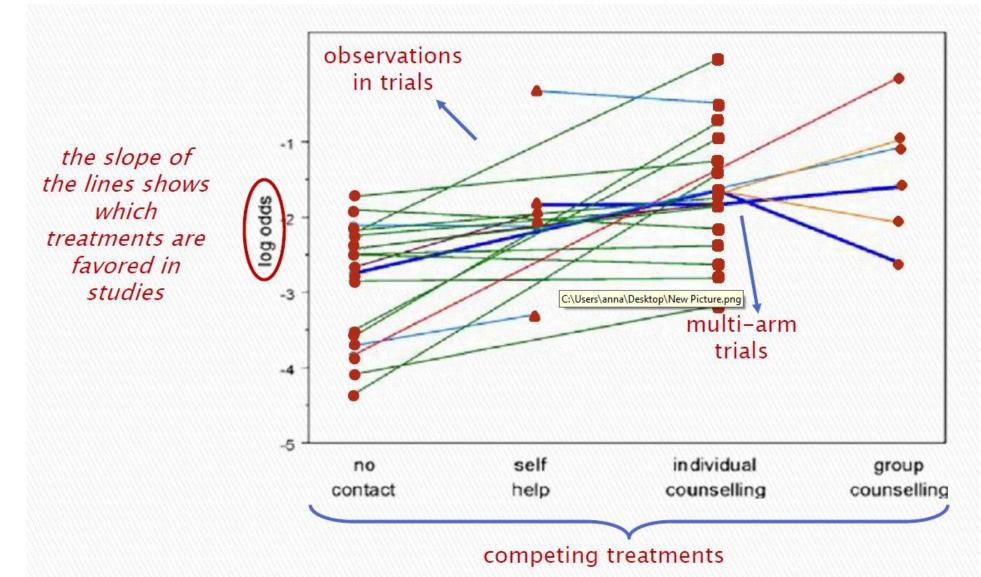
[Examples in Chaimani et al. 2013]

treatments in the respective row and column CTX>TMP/SMX TMP/SMX CTX Cefotaxime CTX+cefixime Gentamicin daily Gentamicin tid A/Clav CTX+netilmicin>cefixime CTX+netilmicin>CTX competing Various Cefixime treatments Cefotaxime>cefixime Isepamicin Amikacin Temocillin >A or A/Clav CTX>A/Clav Sulfafurazole Cefepime>TMP/SMX Ceftazidime>TMP/SMX Cefetamet Netilmicin daily Netilmicin tid CTX>ceftibuten

Matrix showing the available direct comparisons in the network

number of trials comparing the

[Example in loannidis 2006]



Graph showing the data provided by the individual studies of the network

[Example in Lu & Ades 2006]

Presenting the results measures of effect

column efficacy treatment in effects for treatment favor the relative SMD<0

HAL	1-40	1 <u>·49</u>	0-81	1-32	1·11	1·16	0-86	1·16	0-93	0-69	0.85	0 <u>-56</u>	0-48
	(0-93 to 2-11)	(1-03 to 2-15)	(0-53 to 1-22)	(0-85 to 2-06)	(075 to 1·66)	(0·63 to 2·14)	(0-46 to 1-60)	(0·73 to 1·86)	(0-59 to 1-49)	(0-36 to 1-36)	(0.62 to 1.15)	(0-34 to 0-93)	(0-16 to 1-44)
-0-06	RIS	1.06	0-58	0-94	0-80	0-83	0-62	0-83	0-67	0-50	0-61	0-40	0-34
(-0-22 to 0-11)		(0.72 to 1.56)	(0-37 to 0-88)	(0-60 to 1-47)	(0-51 to 1-25)	(0-44 to 1-57)	(0-33 to 1-16)	(0-51 to 1-34)	(0-41 to 1-10)	(0-25 to 0-98)	(0-44 to 0-83)	(0-24 to 0-68)	(0-11 to 1-03)
-0·12	-0-07	OLZ	0-54	0.88	0-75	0.78	0-58	0-78	0.63	0·47	0-57	0-38	0-32
(-0·28 to 0·02)	(-0-22 to 0-08)		(0-37 to 0-79)	(0.58 to 1.36)	(0-49 to 1-13)	(0.43 to 1.44)	(0-33 to 1-00)	(0-52 to 1-17)	(0.40 to 1.00)	(0·24 to 0·89)	(0-44 to 0-74)	(0-23 to 0-61)	(0-11 to 0-95
<u>-0·19</u>	-0·13	-0·06	LIT	1.63	1-38	1-44	1-07	1·44	1·15	0-86	1-05	0·70	0-60
-0·36 to -0·01)	(-0·30 to 0·04)	(-0·22 to 0·10)		(1.06 to 2.54)	(0 91 to 2-12)	(0-81 to 2-60)	(0-57 to 2-00)	(0·92 to 2·28)	(0·71 to 1·91)	(0-47 to 1-59)	(0-78 to 1-43)	(0·44 to 1·11)	(0-20 to 1-77)
<u>-0·19</u>	-0·13	-0-07	-0·01	QTP	0.85	0-88	0.66	0.88	0-71	0-53	0-64	0·43	0-36
-0·37 to -0·01)	(-0·31 to 0·04)	(-0-24 to 0-11)	(-0·18 to 0·17)		(0.52 to 1.35)	(0-46 to 1-70)	(0.34 to 1.25)	(0.53 to 1.46)	(0-42 to 1-20)	(0-27 to 1-05)	(0-45 to 0-91)	(0·25 to 0·73)	(0-12 to 1-10)
<u>-0·19</u>	-0·13	-0.06	-0·01	0.00	ARI	1·04	0-77	1.05	0.84	0.62	0.76	0-50	0-43
-0·36 to -0·02)	(-0·31 to 0·05)	(-0.23 to 0.11)	(-0·18 to 0·17)	(-0.19 to 0.20)		(0·55 to 1·98)	(0-41 to 1-47)	(0.64 to 1.70)	(0.51 to 1.39	(0.32 to 1.24)	(0.55 to 1.06)	(0-30 to 0-85)	(0-14 to 1-29)
<u>-0·20</u>	-0·14	-0.08	-0-02	-0·01	-0.01	CBZ	0-74	1·00	0.80	0.60	0-73	0-48	0-41
-0·36 to -0·01)	(-0·42 to 0·12)	(-0.34 to 0.18)	(-0-28 to 0-24)	(-0·30 to 0·26)	(-0.29 to 0.26)		(0-34 to 1-62)	(0·52 to 1·91)	(0.41 to 1.59)	(0.27 to 1.33)	(0-42 to 1-28)	(0-25 to 0-96)	(0-13 to 1-37
<u>-0.26</u>	-0-20	-0·14	-0.08	-0·07	-0·07	-0·06	ASE	1·35	1-08	0.81	0-98	0-65	0-56
(-0.52 to -0.01)	(-0-46 to 0-05)	(-0·36 to 0·10)	(-0.41 to 0.27)	(-0·34 to 0·20)	(-0·34 to 0·20)	(-0·39 to 0·28)		(0·71 to 2·58)	(0-56 to 2-14)	(0.36 to 1.83)	(0-57 to 1-72)	(0-33 to 1-30)	(0-17 to 1-82
-0·36	-0-30	-0-23	-0·10	-0-17	-0-17	-0·15	-0·10	VAL	0.80	0.60	0-73	0-48	0-41
-0·56 to -0·15)	(-0-50 to -0-10)	(-0-40 to -0-06)	(-0·41 to 0·23)	(-0-38 to 0-05)	(-0-38 to 0-05)	(-0·44 to 0·13)	(-0·37 to 0·18)		(0.47 to 1.37)	(0.30 to 1.20)	(0-51 to 1-05)	(0-28 to 0-83)	(0-13 to 1-25
<u>-0·36</u>	-0-31	-0-24	-0·15	-0-17	-0·18	-0·16	-0·10	-0-01	ZIP	0-75	0·91	0.61	0-52
-0·56 to -0·15)	(-0-51 to -0-10)	(-0-43 to -0-03)	(-0·44 to 0·16)	(-0-39 to 0-05)	(-0·39 to 0·04)	(-0·45 to 0·14)	(-0·39 to 0·18)	(-0-24 to 0-23)		(0-37 to 1-51)	(0·61 to 1·34)	(034 to 106)	(0 17 to 1 58
<u>-0-48</u>	-0·43	<u>-0-36</u>	-0-32	-0·29	-0·29	-0-28	-0-22	-0·13	-0-12	LAM	1-22	0-81	0-69
-0-77 to -0-19)	(-0·71 to -0·14)	(-0-64 to -0-08)	(-0-67 to 0-06)	(-0·58 to 0·00)	(-0·58 to 0·00)	(-0-63 to 0-08)	(-0-57 to 0-12)	(-0·43 to 0·18)	(-0-43 to 0-19)		(0-67 to 2-21)	(0-40 to 1-65)	(0-21 to 2-30
-0·56	-0-50	-0-43	-0-37	-0-37	-0·37	-0-36	-0-30	-0-20	<u>-0.20</u>	-0-08	РВО	0-66	0-57
-0·69 to -0·43)	(-0-63 to -0-38)	(-0-54 to -0-32)	(-0-63 to -0-11)	(-0-51 to -0-23)	(-0·51 to -0·23)	(-0-60 to -0-11)	(-0-53 to -0-07)	(-0-37 to -0-04)	(-0.37 to -0.03)	(-0-34 to 0-18)		(0-44 to 1-00)	(0-20 to 1-62
-0.63	-0-58	-0-51	-0·45	-0·44	-0-45	-0·43	-0-38	-0.28	-0-27	-0·15	-0-07	ТОР	0-85
-0.84 to -0.43)	(-0-78 to -0-37)	(-0-70 to -0-31)	(-0·75 to -0·14)	(-0·66 to -0·23)	(-0-66 to -0-23)	(-0·72 to -0·14)	(-0-66 to -0-09)	(-0.52 to -0.04)	(-0-51 to -0-04)	(-0·46 to 0·15)	(-0-24 to 0-09)		(0-28 to 2-63
-0-88	<u>-0.83</u>	<u>-0-76</u>	<u>-0.70</u>	-0-69	-0-69	-0-68	-0-62	-0-53	-0-52	-0·40	-0·32	-0·25	GBT
(-1-40 to -0-36)	(-1.34 to -0.31)	(-1-27 to -0-24)	(-1.21 to -0.18)	(-1-21 to -0-17)	(-1-21 to -0-17)	(-1-23 to -0-12)	(-1-17 to -0-07)	(-1-05 to 0-01)	(-1-05 to 0-01)	(-0·96 to 0·16)	(-0·82 to 0·.8)	(-0·77 to 0·28)	

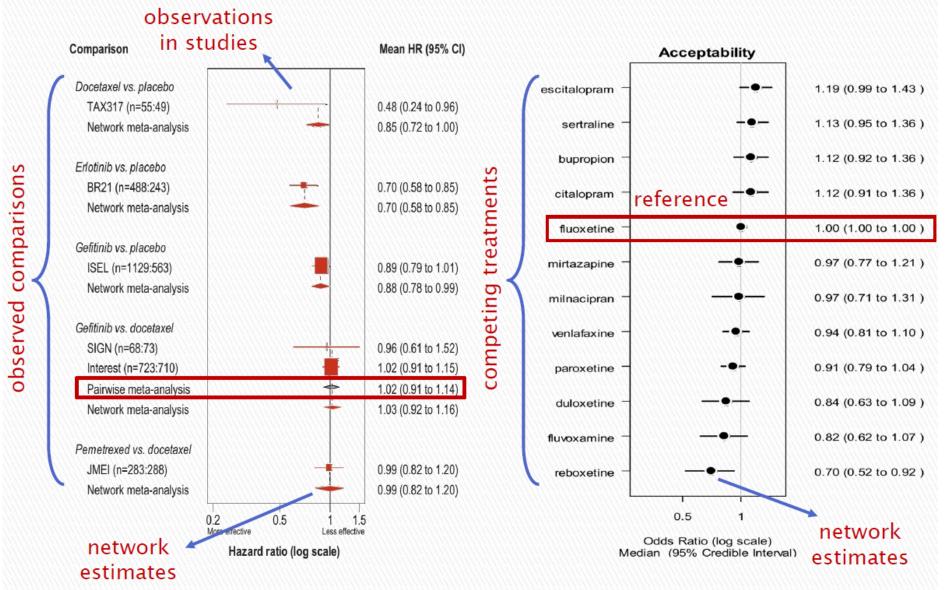
significant effects are in bold and competing treatments underscored font

Table showing all the pairwise relative treatment effects with their 95% CI for one or two outcomes

[Example in Cipriani et al. 2011]

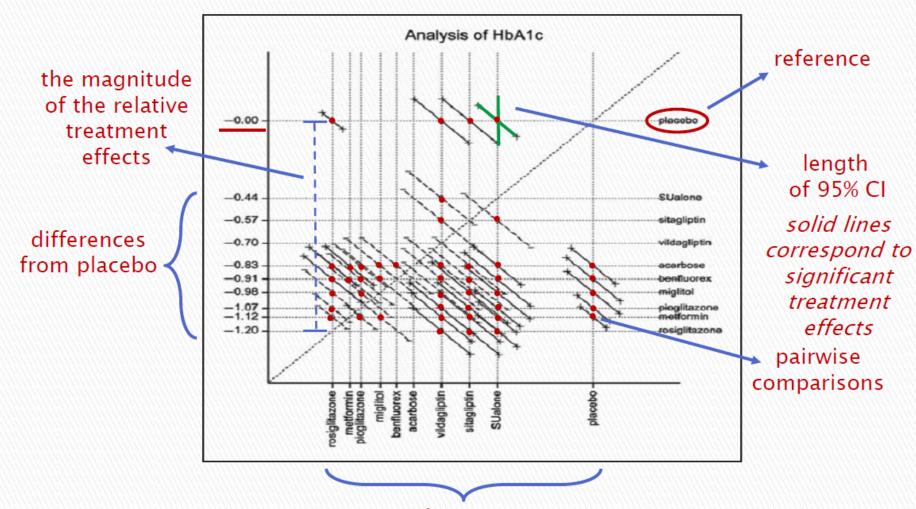
effects

relative



Forest plot with the treatment effects estimates for the pairwise comparisons

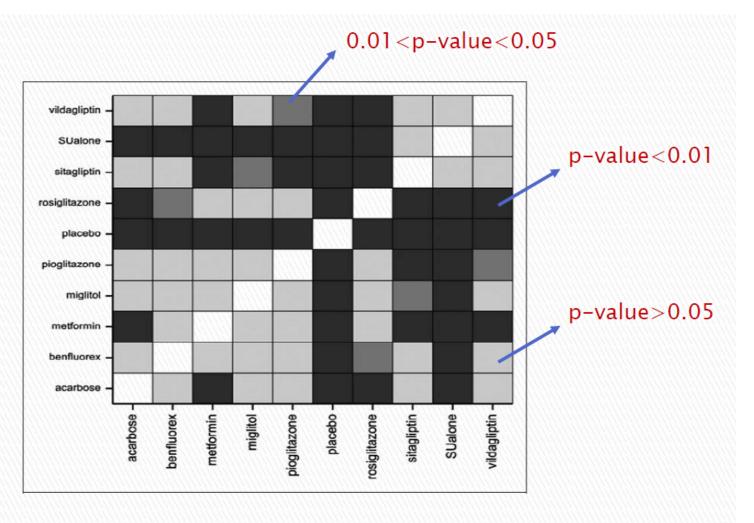
[Examples in Hawkins et al. 2009 & Hoaglin et al. 2011]



competing treatments

'Hsu mean-mean plot' showing the network estimates with the 95% CI for all pairwise comparisons

[Example in Senn et al. 2013]



Shade plot showing the p-values of the treatment effects for all pairwise comparisons in the network

[Example in Senn et al. 2013]

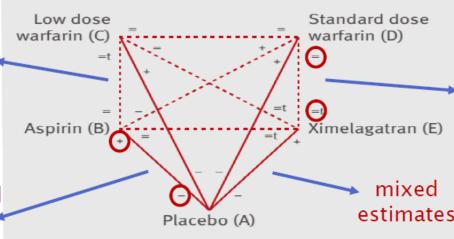
indirect estimates

in comparisons with significant relative treatment effects (based on the network estimates):

'+' the favored treatment

'-' the non-favored

treatment



- More effective
- Less effective
- = No difference
- t Treatment favoured by trend when no difference in comparison

---- RCT

----- Indirect comparison

Comparison (relative risk (95% confidence interval))

B v A: 0.64 (0.44 to 0.88) D v B: 0.55 (0.32 to 0.92)
C v A: 0.35 (0.19 to 0.60) E v B: 0.53 (0.26 to 1.07)
D v A: 0.35 (0.24 to 0.52) D v C: 1.00 (0.50 to 1.99)
E v A: 0.34 (0.18 to 0.61) E v C: 0.97 (0.42 to 2.25)
C v B: 0.55 (0.28 to 1.07) E v D: 0.97 (0.47 to 2.00)

in comparisons with non-significant relative treatment effects (based on the network estimates): '=t' the favored

treatment

estimates '=' the non-favored treatment

Network graph presenting the relative treatment effects for each pairwise comparison

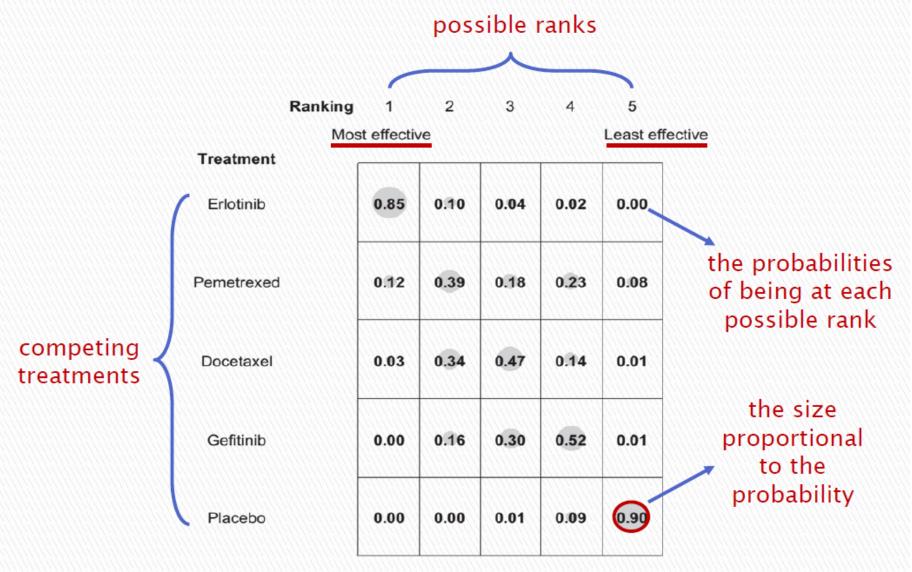
[Example in Fadda et al. 2011]

Presenting the results ranking

Using probability of being the best

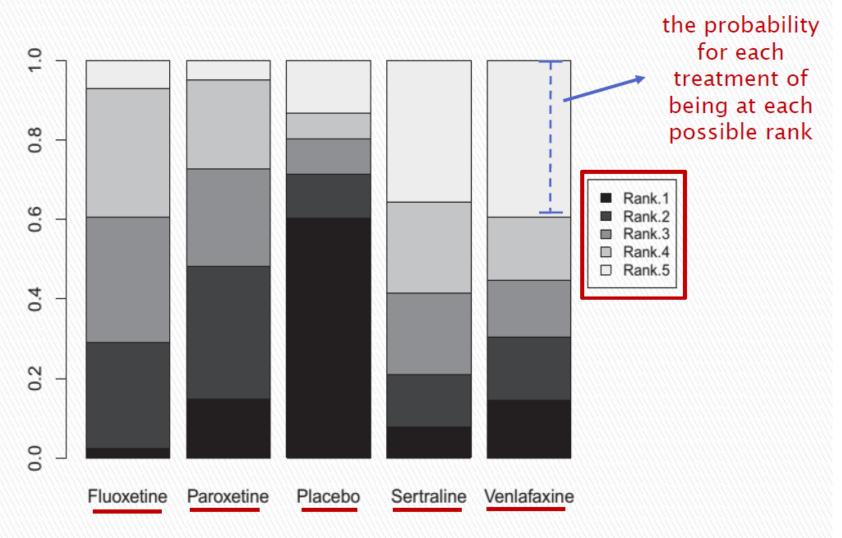
Using probabilities of being at each possible rank

Using SUCRAS



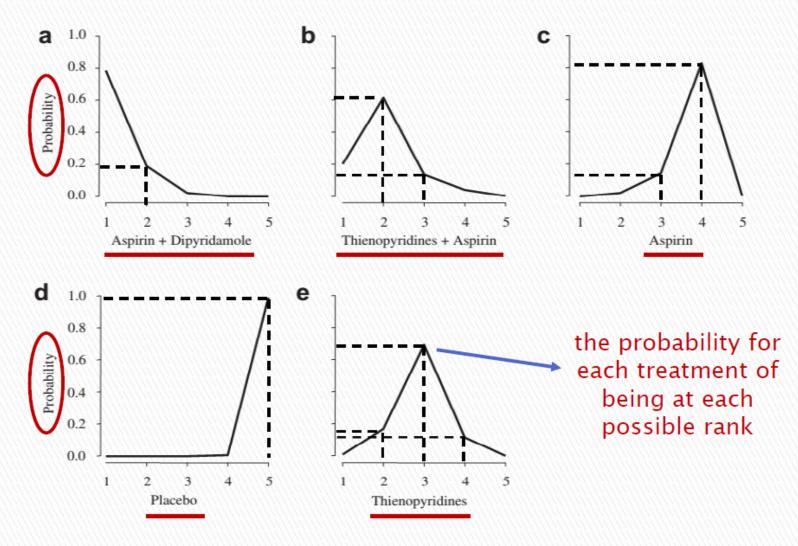
'Bubble-plot' including the ranking probabilities for all treatments

[Example in Hawkins et al. 2009]



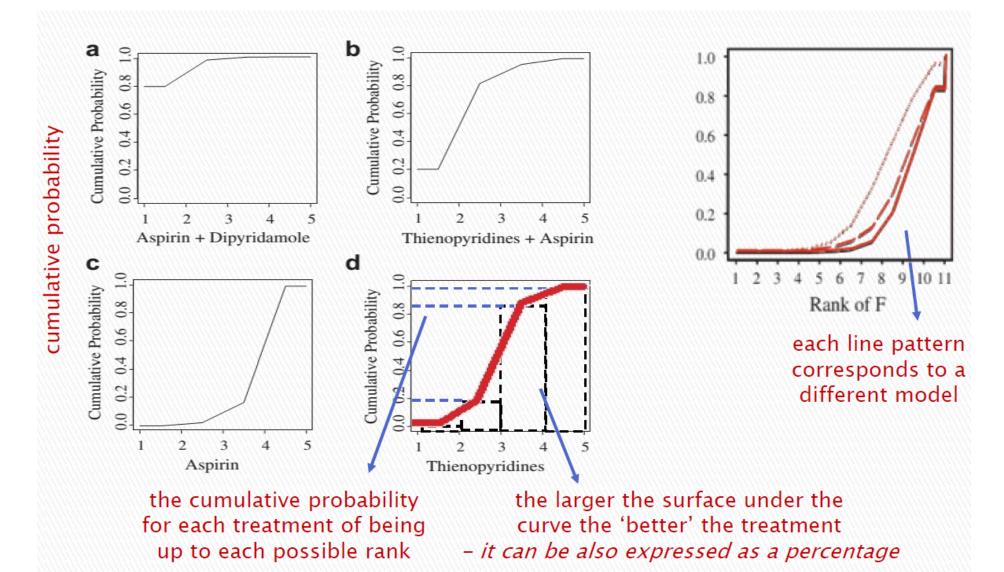
Bar plots showing the probability for each treatment of being at a specific rank

[Example in van Valkenhoef et al. 2012]



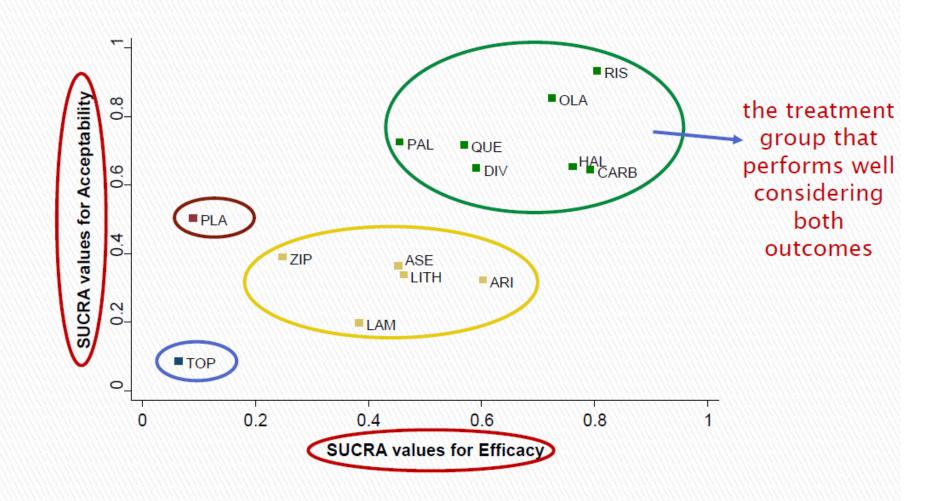
'Rankograms' showing the probability for each treatment of being at a specific rank

[Example in Salanti et al. 2011]



'SUCRA plots' showing the cumulative probability for each treatment of being up to a specific rank

[Examples in Salanti et al. 2011 & Salanti et al. 2010]



Scatterplot showing jointly the ranking results for two different outcomes

[Example in Chaimani et al. 2013]

HAL	.47 2:11)	1-49	0-81	1-32	1-11	1-16	0-86	1-16	0-93	0-69	0-85	0.56	0-48
0.95/0		(1-03 to 2-15)	(0-53 to 1-22)	(0-85 to 2-06)	(0-75 to 1-66)	(0-63 to 2-14)	(0-46 to 1-60)	(0-73 to 1-86)	(0-59 to 1-49)	(0-36 to 1-36)	(0-62 to 1-15)	(0-34 to 0-93)	(0-16 to 1-44)
-0-06	RIS	78 (1-56)	0-58	0-94	0-80	0-83	0-62	0-83	0-67	0-50	0-61	0-40	0-34
-0-22 to 0-11)	0.94/0.		(0-37 to 0-88)	(0-60 to 1-47)	(0-51 to 1-25)	(0-44 to 1-57)	(0-33 to 1-16)	(0-51 to 1-34)	(0-41 to 1-10)	(0-25 to 0-98)	(0-44 to 0-83)	(0-24 to 0-68)	(0-11 to 1-03)
-0·12	-0-07	OLZ	5 <u>4</u>	0-88	0-75	0-78	0-58	0-78	0-63	0-47	0-57	0-38	0-32
-0·28 to 0·02)	(-0-22 to 0-08)	0.78/0.	81 <u>00-79)</u>	(0-58 to 1-36)	(0-49 to 1-13)	(0-43 to 1-44)	(0-33 to 1-00)	(0-52 to 1-17)	(0-40 to 1-00)	(0-24 to 0-89)	(0-44 to 0-74)	(0-23 to 0-61)	(0-11 to 0-95
<u>-0·19</u>	-0·13	-0.06	LIT		1-38	1-44	1-07	1-44	1·15	0-86	1-05	0-70	0-60
0·36 to -0·01)	(-0·30 to 0·04)	(-0.22 to 0.10	0.64/0.2		(0-91 to 2-12)	(0-81 to 2-60)	(0-57 to 2-00)	(0-92 to 2-28)	(0·71 to 1·91)	(0-47 to 1-59)	(0-78 to 1-43)	(0-44 to 1-11)	(0-20 to 1-77)
-0·19	-0·13	-0-07	-0-01	QTP		0-88	0-66	0-88	0-71	0-53	0-64	0-43	0-36
0·37 to -0·01)	(-0·31 to 0·04)	(-0-24 to 0-11)	(-0-18 to 0-17	0.64/0.7		(0-46 to 1-70)	(0-34 to 1-25)	(0-53 to 1-46)	(0-42 to 1-20)	(0-27 to 1-05)	(0-45 to 0-91)	(0-25 to 0-73)	(0-12 to 1-10)
-0·19	-0·13	-0.06	-0-01	0-00	ARI	7 to 1-98)	0-77	1-05	0.84	0-62	0-76	0-50	0-43
0·36 to -0·02)	(-0·31 to 0·05)	(-0.23 to 0.11)	(-0-18 to 0-17)	(-0-19 to 0-2	0.61/0.5		(0-41 to 1-47)	(0-64 to 1-70)	(0-51 to 1-39	(0-32 to 1-24)	(0-55 to 1-06)	(0-30 to 0-85)	(0-14 to 1-29)
<u>-0·20</u>	-0-14	-0.08	-0-02	-0-01	-0-01	CBZ		1-00	0.80	0-60	0-73	0-48	0-41
0·36 to -0·01)	(-0-42 to 0-12)	(-0.34 to 0.18)	(-0-28 to 0-24)	(-0-30 to 0-26)	(-0-29 to 0-3	0.60/0.60		(0-52 to 1-91)	(0.41 to 1.59)	(0-27 to 1-33)	(0-42 to 1-28)	(0-25 to 0-96)	(0-13 to 1-37)
<u>-0-26</u>	-0-20	-0·14	-0.08	-0-07	-0-07	-0-06	ASE		1-08	0-81	0-98	0-65	0-56
0-52 to -0-01)	(-0-46 to 0-05)	(-0·36 to 0·10)	(-0.41 to 0.27)	(-0-34 to 0-20)	(-0-34 to 0-20)	(-0-39 to 0	0.55/0.36		(0-56 to 2-14)	(0-36 to 1-83)	(0-57 to 1-72)	(0-33 to 1-30)	(0-17 to 1-82)
-0·36	- <u>0-30</u>	-0-23	-0-10	-0-17	-0-17	-0·15	-0-10	VAL		0-60	0-73	0-48	0-41
0·56 to -0·15)	(-0-50 to -0-10)	(-0-40 to -0-06)	(-0-41 to 0-23)	(-0-38 to 0-05)	(-0-38 to 0-05)	(-0·44 to 0·13)	(-0-37 to 0	0.50/0.48		(0-30 to 1-20)	(0-51 to 1-05)	(0-28 to 0-83)	(0-13 to 1-25)
<u>-0-36</u>	<u>-0-31</u>	<u>-0-24</u>	-0·15	-0-17	-0·18	-0-16	-0-10	-0-0	ZIP	0-75	0-91	0.61	0-52
0-56 to -0-15)	(-0-51 to -0-10)	(-0-43 to -0-03)	(-0·44 to 0·16)	(-0-39 to 0-05)	(-0·39 to 0·04)	(-0-45 to 0-14)	(-0-39 to 0-18)	(-0-24 to 0	.47/0.41	0-37 to 1-51)	(0-61 to 1-34)	(0 34 to 106)	(0 17 to 1 58)
<u>-0-48</u>	<u>-0-43</u>	<u>-0-36</u>	-0-32	-0-29	-0-29	-0-28	-0-22	-0·13	-0-1	LAM	1-22	0-81	0-69
0-77 to -0-19)	(-0-71 to -0-14)	(-0-64 to -0-08)	(-0-67 to 0-06)	(-0-58 to 0-00)	(-0-58 to 0-00)	(-0-63 to 0-08)	(-0-57 to 0-12)	(-0·43 to 0·18)	(-0-43 to 0	.40/0.21	0-67 to 2-21)	(0-40 to 1-65)	(0-21 to 2-30)
<u>-0-56</u>	<u>-0-50</u>	<u>-0-43</u>	-0-37	=0-37	<u>-0-37</u>	-0-36	-0-30	-0-20	<u>-0.20</u>	-0-0	PBO	0-66	0-57
0-69 to -0-43)	(-0-63 to -0-38)	(-0-54 to -0-32)	(-0-63 to -0-11)	(=0-51 to =0-23)	(-0-51 to -0-23)	(-0-60 to -0-11)	(-0-53 to -0-07)	(-0-37 to -0-04)	(-0-37 to -0-03)	(-0-34 to 0)	.36/0.30	0-44 to 1-00)	(0-20 to 1-62)
-0-63	-0-58	-0-51	=0:45	-0-44	-0-45	-0-43	-0-38	<u>-0-28</u>	-0-27	-0·15	-0-	TOP	0-85
0-84 to -0-43)	(-0-78 to -0-37)	(-0-70 to -0-31)	(-0:75 to -0:14)	(-0-66 to -0-23)	(-0-66 to -0-23)	(-0-72 to -0-14)	(-0-66 to -0-09)	(-0-52 to -0-04)	(-0-51 to -0-04)	(-0·46 to 0·15)	(-0-241 0	23/0.09	(0-28 to 2-63)
-0-88	-0-83	-0-76	-0.70	-0-69	-0-69	-0-68	-0-62	-0-53	-0-52	-0-40	-0-32	-c	GBT 3/0.12
1-40 to -0-36)	(-1-34to-0-31)	(-1-27 to -0-24)	(-1.21to-0.18)	(-1-21 to -0-17)	(-1-21 to -0-17)	(-1-23 to -0-12)	(-1-17 to -0-07)	(-1-05 to 0-01)	(-1-05 to 0-01)	(-0-96 to 0-16)	(-0-82 to 0-18)	(-0-77 O. 1	

competing treatments ordered according to their relative ranking for efficacy

Table showing all the pairwise relative treatment effects with their 95% CI for one or two outcomes along with the SUCRA values

Erythropoiesis-stimulating agents for anaemia in adults with chronic kidney disease: a network meta-analysis (Review)

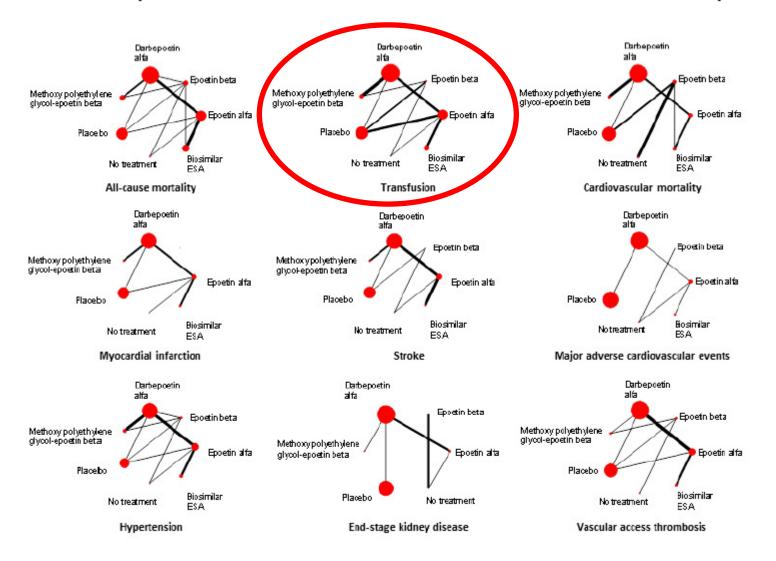
Palmer SC, Saglimbene V, Mavridis D, Salanti G, Craig JC, Tonelli M, Wiebe N, Strippoli GFM

Copyright © 2014 The Cochrane Collaboration. Published by JohnWiley & Sons, Ltd.

OBJECTIVES

To compare the efficacy and safety of ESAs (epoetin alfa, epoetin beta, darbepoetin alfa, or methoxy polyethylene glycol-epoetin beta, and biosimilar ESAs, against each other, placebo, or no treatment) to treat anaemia in adults with CKD.

Figure 5. Networks of the treatment efficacy and safety of ESA drugs in the treatment of anaemia in chronic kidney disease. Values lower than I favour the active treatment in the comparison



(... Continued) Assessment of clinical and methodological Study or subgroup Intervention Comparator Odds Ratio Weight Odds Ratio H,Random,95% CI H,Random,95% n/N Analysis I.I. Comparison I ESA versus ESA or placebo/no treatment, Outcom Subtotal (95% CI) 2012 2026 100.0 % 0.53 [0.46, 0.63] Total events: 297 (Intervention), 496 (Comparator) Review: Erythropoiesis-stimulating agents for anaemia in adults with chronic kidney disease: a network meta-analysis Heterogeneity: not applicable Test for overall effect: Z = 7.71 (P < 0.00001) Comparison: I ESA versus ESA or placebo/no treatment 4 Epoetin alfa versus control Outcome: I Blood transfusion Patel 2012 4/118 0/39 100.0 % 3.10 [0.16, 58.97] Subtotal (95% CI) 118 39 100.0 % 3.10 [0.16, 58.97] Study or subgroup Intervention Comparator Odds Ratio Total events: 4 (Intervention), 0 (Comparator) Heterogeneity: not applicable H,Random,95% n/N n/N Test for overall effect: Z = 0.75 (P = 0.45) 5 Epoetín beta versus no control I Epoetín alfa versus placebo Van Biesen 2005 2/22 100.0 % 0.35 [0.06, 2.18] 4/18 Kleinman 1989 3/7 24.9 0/7 Subtotal (95% CI) 22 18 100.0 % 0.35 [0.06, 2.18] Canadían EPO Study 1990 2/67 23/32 36.4 Total events: 2 (Intervention), 4 (Comparator) Heterogeneity: not applicable Roth 1994 4/43 9/40 38.7 Test for overall effect: Z = 1.12 (P = 0.26) Subtotal (95% CI) 117 79 100.0 9 6 Epoetin alfa versus darbepoetin alfa Total events: 6 (Intervention), 35 (Comparator) Akizawa 2011 0/160 1/161 2.8 % 0.33 [0.01, 8.24] Heterogeneity: $Tau^2 = 3.61$; $Chi^2 = 10.53$, df = 2 (P = 0.01); $I^2 = 81\%$ Locatelli 2001 3/37 6/129 14.2 % 1.81 [0.43, 7.61] Test for overall effect: Z = 2.10 (P = 0.036) Nissenson 2002 37/335 17/369 82.9 % 2.57 [1.42, 4.66] 2 Epoetín beta versus placebo Subtotal (95% CI) 532 659 100.0 % Bennett 1991 0/90 1/41 2.31 [1.34, 3.97] Total events: 40 (Intervention), 24 (Comparator) Bahlmann 1991 5/53 28/46 89.6 Heterogeneity: $Tau^2 = 0.0$; $Chi^2 = 1.64$, df = 2 (P = 0.44); $I^2 = 0.0\%$ Test for overall effect: Z = 3.03 (P = 0.0025) Subtotal (95% CI) 143 100.0 9 7 Epoetín alfa versus biosímilar ESA Total events: 5 (Intervention), 29 (Comparator) Krivoshiev 2010 1/230 3/232 5.5 % 0.33 [0.03, 3.23] Heterogeneity: $Tau^2 = 0.0$; $Chi^2 = 0.21$, df = 1 (P = 0.64); $I^2 = 0.0\%$ Krivoshiev 2008 33.6 % 0.90 [0.36, 2.25] 9/304 10/305 Test for overall effect: Z = 4.95 (P < 0.00001) 3 Darbepoetin alfa versus placebo Martin 2007 11/192 46/560 60.9 % 0.68 [0.34, 1.34] TREAT Study 2005 297/2012 496/2026 100.0 Subtotal (95% CI) 726 1097 100.0 % 0.72 [0.42, 1.22] Total events: 21 (Intervention), 59 (Comparator) 0.001 0.01 0.1 1 10 100 1000 Heterogeneity: Tau2 = 0.0; Chi2 = 0.70, df = 2 (P = 0.70); I2 = 0.0% Favours intervention Favours comparator Test for overall effect: Z = 1.22 (P = 0.22) 8 Epoetin beta versus methoxy polyethylene glycol-epoetin beta 100.0 % 0.83 [0.17, 4.15] AMICUS Study 2007 7/135 Calanda (OSO/ CD) 46 135 100.0 % 0.83 [0.17, 4.15] rator) (... Continued) Odds Ratio Odds Ratio Study or subgroup Comparator Weight Intervention M-H,Random,95% M-H,Random,95% ethylene glycol-epoetin beta 0/25 4.7 % 0.59 [0.02, 15.14] n/N n/N 2.88 [0.90, 9.23] ARCTOS Study 2008 11/162 4/162 23.7 % 0.001 0.01 0.1 1 10 100 1000 CORDATUS Study 2011 0.48 [0.16, 1.44] 5/154 10/153 25.4 % Favours intervention Favours comparator PATRONUS Study 2010 32/244 39/245 46.2 % 0.80 [0.48, 1.32] Subtotal (95% CI) 606 100.0 % 0.94 [0.45, 1.95] Total events: 48 (Intervention), 54 (Comparator)

0.001 0.01 0.1 1 10 100 1000

Favours intervention Favours comparator

Heterogeneity: Tau2 = 0.23; Chi2 = 5.35, df = 3 (P = 0.15); I2 = 44%

Test for overall effect: Z = 0.18 (P = 0.86)

Assessment of similarity (transitivity) across treatment comparisons

Evaluation of the assumption is important and its plausibility determines the validity of the network meta-analysis results.

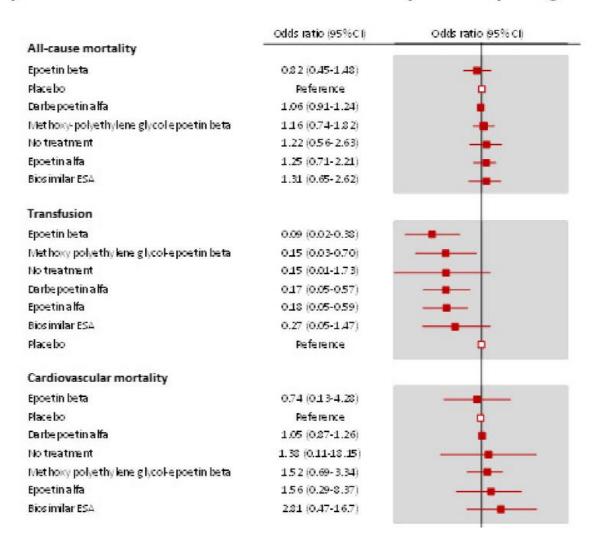
We inferred about the assumption of transitivity:

- 1. We assessed whether the included interventions were similar when they were evaluated in studies with different designs, for example, whether ESAs are administered the same way in studies comparing ESAs to placebo and in those comparing ESAs to other ESAs
- 2. We compared the distribution of the potential effect modifiers (age, stage of CKD, duration of treatment) across the different pairwise comparisons.

The inconsistency factor is the absolute difference in the log odds ratio estimated from indirect and direct treatment comparisons and is reported together with the 95% confidence interval. A 95% confidence interval that includes zero indicates that the result is compatible with zero inconsistency between effect estimates using indirect (networkmeta-analysis) and direct (conventional pairwise meta-analysis) treatment comparisons.

Transfusion		
Epoetin alfa - epoetin beta - placebo – no treatment	2.09	0.00-6.91
Epoetin alfa - darbepoetin alfa - placebo	1.97	0.00-4.20
Epoetin beta - darbepoetin alfa – methoxy polyethylene glycol-epoetin beta - placebo	1.26	0.00-3.39

Figure 6. Forest plots for results from network meta-analyses comparing ESAs versus placebo



Annals of Internal Medicine RESEARCH AND REPORTING METHODS

The PRISMA Extension Statement for Reporting of Systematic Reviews Incorporating Network Meta-analyses of Health Care Interventions: Checklist and Explanations

Brian Hutton, PhD, MSc; Georgia Salanti, PhD; Deborah M. Caldwell, PhD, MA, BA; Anna Chaimani, PhD; Christopher H. Schmid, PhD; Chris Cameron, MSc; John P.A. Ioannidis, MD, DSc; Sharon Straus, MD, MSc; Kristian Thorlund, PhD; Jeroen P. Jansen, PhD; Cynthia Mulrow, MD, MSc; Ferrán Catalá-López, PhD, MPH, PharmD; Peter C. Gøtzsche, MD, MSc; Kay Dickersin, PhD, MA; Isabelle Boutron, MD, PhD; Douglas G. Altman, DSc; and David Moher, PhD

The PRISMA statement is a reporting guideline designed to improve the completeness of reporting of systematic reviews and meta-analyses. Authors have used this guideline worldwide to prepare their reviews for publication. In the past, these reports typically compared 2 treatment alternatives. With the evolution of systematic reviews that compare multiple treatments, some of them only indirectly, authors face novel challenges for conducting and reporting their reviews. This extension of the PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-analyses) statement was developed specifically to improve the reporting of systematic reviews incorporating network meta-analyses.

A group of experts participated in a systematic review, Delphi survey, and face-to-face discussion and consensus meeting to establish new checklist items for this extension statement. Cur-

rent PRISMA items were also clarified. A modified, 32-item PRISMA extension checklist was developed to address what the group considered to be immediately relevant to the reporting of network meta-analyses.

This document presents the extension and provides examples of good reporting, as well as elaborations regarding the rationale for new checklist items and the modification of previously existing items from the PRISMA statement. It also highlights educational information related to key considerations in the practice of network meta-analysis. The target audience includes authors and readers of network meta-analyses, as well as journal editors and peer reviewers.

Ann Intern Med. 2015;162:777-784. doi:10.7326/M14-2385 www.annals.org For author affiliations, see end of text.

Table. Checklist of Items to Include When Reporting a System RESEARCH AND REPORTING METHODS

PRISMA Extension for Network Meta-analysis

Section/Topic	Item # *	Checklist Item†					
TITLE			Table-Continued				
Title	1	Identify the report as a systemati meta-analysis).	140/E-Continued				
ABSTRACT		mesa-unuyarap.	Section/Topic	Item # *	Checklist Item†	Reported on Page #	
Structured summary	2	Provide a structured summary in Background: main objectives	RESULTS±				
		Methods: data sources; study and synthesis methods, such Results: number of studies and	Study selection	17	Give numbers of studies screened, assessed for eligibility, and included in the review, with reasons for exclusions at each stage, ideally with a flow diagram.		
	confidence/c to summarize		Presentation of network structure	53	Provide a network graph of the included studies to enable visualization of the geometry of the treatment network.		
INTRODUCTION		brevity. Discussion/Conclusions: limita Other: primary source of fundi	Summary of network geometry	S4	Provide a brief overview of characteristics of the treatment network. This may include commentary on the abundance of trials and randomized patients for the different interventions and pairwise comparisons in the network, gaps of evidence in the treatment network, and potential biases reflected by the network structure.		
Rationale	3	Describe the rationale for the rev why a network meta-analysis	Study characteristics	18	For each study, present characteristics for which data were extracted (e.g., study size, PICOS,		
Objectives	4	Provide an explicit statement of a interventions, comparisons,	Risk of bias within studies	19	follow-up period) and provide the citations. Present data on risk of bias of each study and, if available, any outcome level assessment.		
METHODS Protocol and registration	5	Indicate whether a review protoc and, if available, provide reg	Results of individual studies	20	For all outcomes considered (benefits or harms), present, for each study: 1) simple summary data for each intervention group, and 2) effect estimates and confidence intervals. Modified		
Eligibility criteria	6	Specify study characteristics (e.g years considered, language, Clearly describe eligible trea have been clustered or merg	Synthesis of results	21	approaches may be needed to deal with information from larger networks. Present results of each meta-analysis done, including confidence/credible intervals. In larger networks, authors may focus on comparisons versus a particular comparator (e.g., placebo or		
Information sources	7	Describe all information sources authors to identify additiona			standard care), with full findings presented in an appendix. League tables and forest plots may be considered to summarize pairwise comparisons. If additional summary measures were		
Search	8	Present full electronic search stra it could be repeated.	Fundametica for	S5	explored (such as treatment rankings), these should also be presented.		
Study selection Data collection process	9	State the process for selecting st and, if applicable, included i Describe method of data extract	Exploration for inconsistency	33	Describe results from investigations of inconsistency. This may include such information as measures of model fit to compare consistency and inconsistency models, P values from statistical tests, or summary of inconsistency estimates from different parts of the treatment		
Data collection process	10	and any processes for obtain			network.		
Data items	11	List and define all variables for w assumptions and simplificati	Risk of bias across studies	22	Present results of any assessment of risk of bias across studies for the evidence base being studied.		
Geometry of the network	S1	Describe methods used to explo potential biases related to it summarized for presentation the evidence base to reader	Results of additional analyses	23	Give results of additional analyses, if done (e.g., sensitivity or subgroup analyses, meta-regression analyses, alternative network geometries studied, alternative choice of prior distributions for Bayesian analyses, and so forth).		
Risk of bias within individual studies	12	Describe methods used for asset whether this was done at the	DISCUSSION		,		
Summary measures	13	in any data synthesis. State the principal summary meas of additional summary meas cumulative ranking curve (SL	Summary of evidence	24	Summarize the main findings, including the strength of evidence for each main outcome; consider their relevance to key groups (e.g., health care providers, researchers, and policymakers).		
Planned methods of analysis	14	summary findings from meta Describe the methods of handlin meta-analysis. This should in Handling of multigroup trials;	Limitations	25	Discuss limitations at study and outcome level (e.g., risk of bias), and at review level (e.g., incomplete retrieval of identified research, reporting bias). Comment on the validity of the assumptions, such as transitivity and consistency. Comment on any concerns regarding network geometry (e.g., avoidance of certain comparisons).		
		Selection of variance structure; Selection of prior distributions Assessment of model fit.	Conclusions	26	Provide a general interpretation of the results in the context of other evidence, and implications for future research.		
Assessment of inconsistency	52	Describe the statistical methods the treatment network(s) stu	FUNDING				
Risk of bias across studies	15	Specify any assessment of risk of bias, selective reporting with Describe methods of additional:	Funding	27	Describe sources of funding for the systematic review and other support (e.g., supply of data); role of funders for the systematic review. This should also include information regarding whether		
Additional analyses	10	include, but not be limited to Sensitivity or subgroup analyst Meta-regression analyses; Alternative formulations of the	ons for Bayesian analyses (if applicat		of funders for the systematic review. This should also include information regarding whether funding has been received from manufacturers of treatments in the network and/or whether some of the authors are content experts with professional conflicts of interest that could affect use of treatments in the network.		

(Continued on following page)



SCUOLA DI METODOLOGIA DELLA RICERCA CLINICA

2024 - 10^a EDIZIONE

MODULI SPECIALISTICI - S3



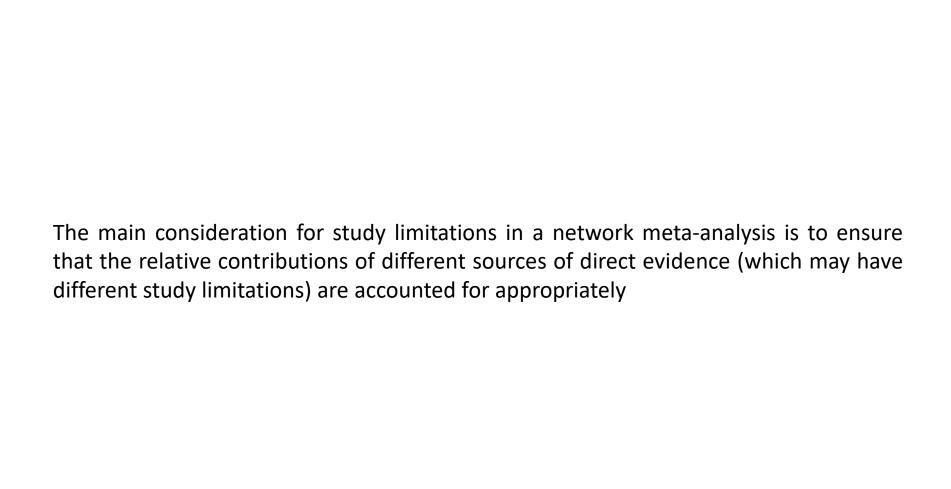
GIOVEDÌ 9 MAGGIO 2024

NEGRAR DI VALPOLICELLA (VR)

Centro Formazione IRCCS "Sacro Cuore - Don Calabria"

La valutazione della certezza delle prove

(M. Cinquini)



Determinants of certainty in a body of evidence

GRADE

- A body of evidence starts as: high | ⊕⊕⊕⊕
- 5 factors that can lower quality
 - 1. Risk of bias criteria
 - 2. Inconsistency (or heterogeneity)
 - 3. Indirectness (PICO and applicability)
 - 4. Imprecision
 - 5. Publication bias



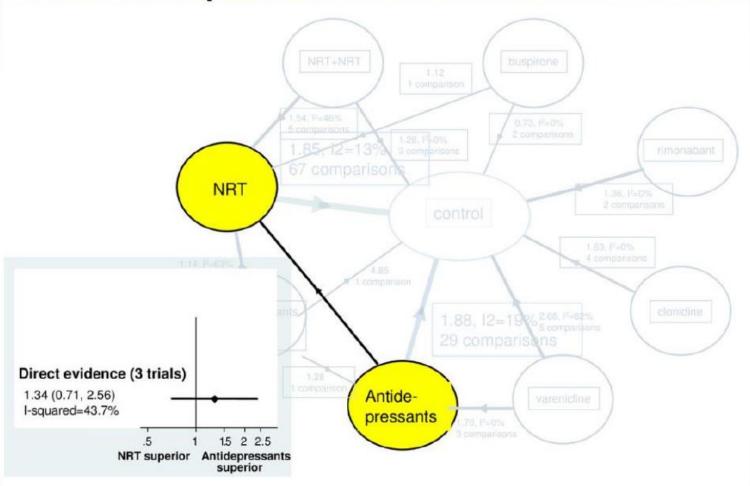


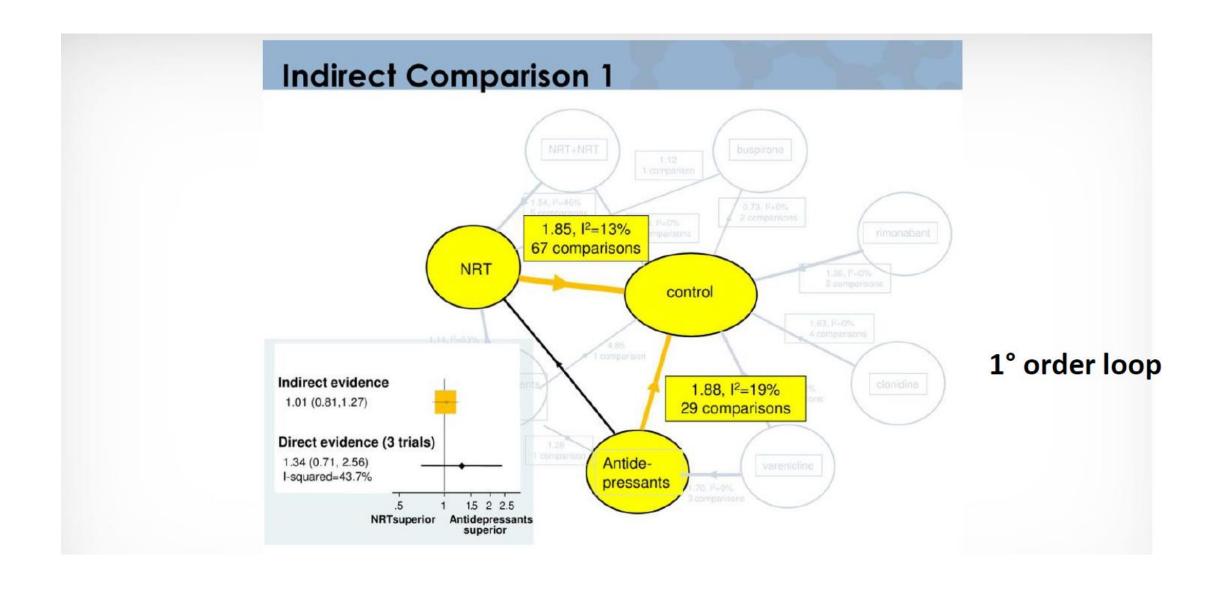




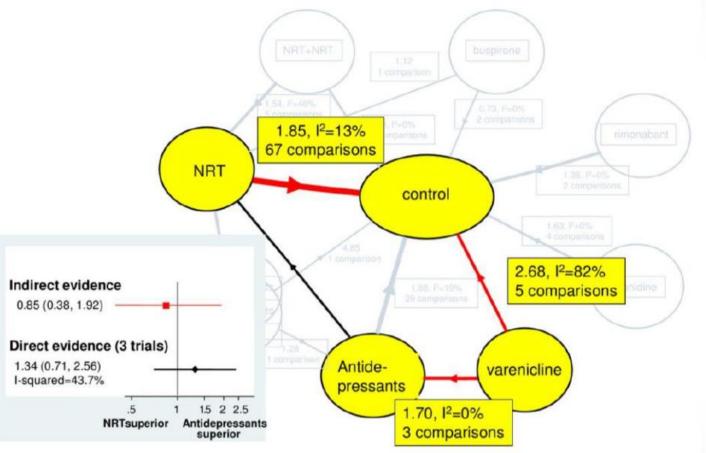


Direct Comparison





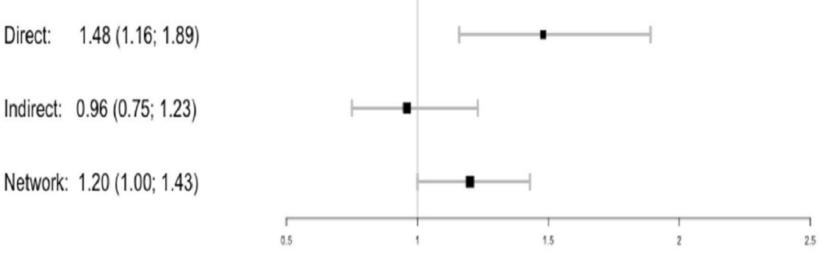
Indirect Comparison 2



2° order loop

Steps for assessing certainty in NMA

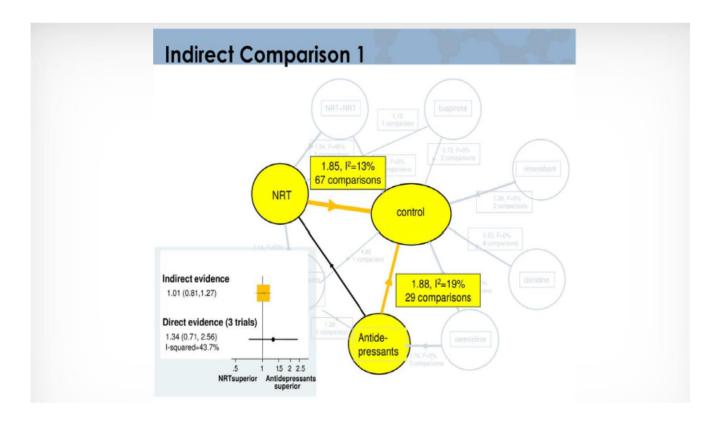
- Step 1: Presenting direct and indirect effect estimates and 95% CI
- Making valid inferences on the basis of a NMA (valid judgment about certailty) requires understanding how much direct and indirect evidence contribute to the NMA effect estimates
- Node splitting approach which separates evidence on a particular comparison (a "node") into direct and indirect estimates of treatment effect



Steps for assessing certainty in NMA

- Step 2: Rating of quality of direct and indirect effect estimates for all domains except imprecision
- To keep the quality rating of the indirect evidence manageable,

we suggest a **focus on first order loops**, which usually contribute
most information to the indirect
estimate.



Steps for assessing certainty in NMA

Presenting and rating of quality of NMA effect estimates

- Steps 3: Use the certainty of evidence of direct or indirect estimate evidence on the basis of which contribute the most to NMA evidence (node splitting) for all domains except imprecision
- Step 4: Assess imprecision of the NMA estimate

Sintesi percorso per valutare certezza evidenza NMA

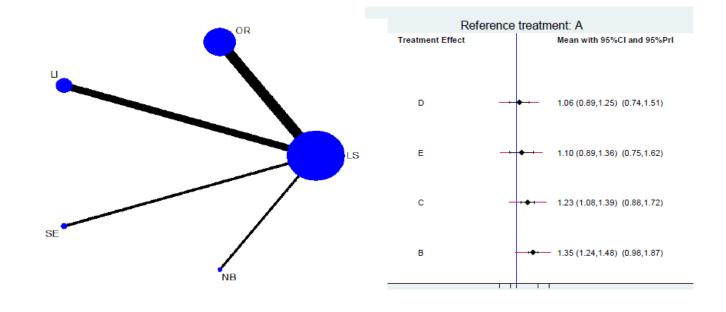
Se c'è solo evidenza diretta

- 1. Valutare certezza evidenza diretta (dalle MA pairwise per tutti i domini tranne imprecision)
- 2. Valutare imprecisione della stima NMA, non pairwise (Il contributo della evidenza che deriva da tutto il network può comunque aumentare la precisione della stima. Questa è una delle ragioni per cui si fa una NMA)

Approccio non contestualizzato: si abbassa per imprecisione se i CI crossano la linea di non effetto

Approccio parzialmente contestualizzato : gli autori della RS devono stabilire a priori le soglie per effetto trivial, piccolo, modesto, grande. Si contano il numero di soglie che vengono attraversate dai CI;

se crossano una soglia si abbassa di un livello, se crossano due soglie si abbassa di due livelli se crossano 3 o + soglie si abbassa di 3 livelli



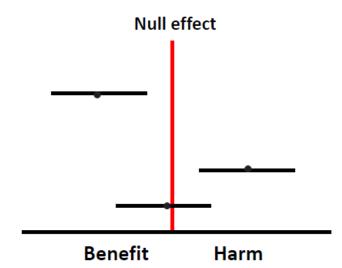
Imprecisione della stima NMA

Non è possibile calcolare OIS perché non si conosce il sample size della NMA

Approccio non contestualizzato:

Non interessa la dimensione dell'effetto; solo la direzione

si abbassa per imprecision se i CI crossano la linea di non effetto

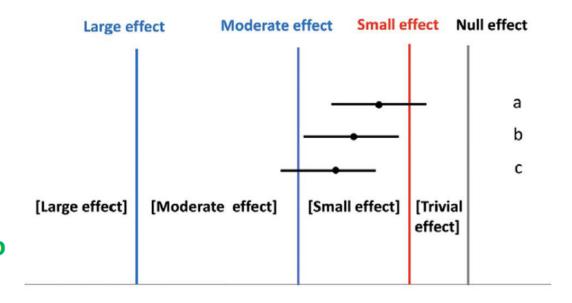


Imprecisione della stima NMA

Approccio Contestualizzato: interessa la dimensione dell'effetto:

irrilevante, piccola, moderata, grande I membri del panel devono stabilire le soglie (valori dell'esito di interesse) per effetto irrilevante, piccolo, moderato, grande, (possibilmente basandosi su dati della letteratura). Si contano il numero di soglie che vengono attraversate dai CI:

se crossano una soglia si abbassa di un livello se crossano due soglie si abbassa di 2 livelli se crossano 3 o + soglie si abbassa di 3 livelli



• Se c'è solo evidenza indiretta

Si considerano solo le due comparison del primo loop (se sono interessato ad B vs C, considero le pairwise di A vs B e di A vs C Si valuta certainty delle due comparison (pairwise) indirette del primo loop per tutte le dimensioni tranne imprecision.

Si considera la certezza più bassa tra le due

Si valuta imprecisione della stima della NMA come sopra



Devo vedere quale delle due certezze contribuisce di più alla stima network

• Se una stima (diretta o indiretta) contribuisce di più alla stima network

Valuto la certezza per tutte le dimensioni tranne imprecision della evidenza che contribuisce di più seguendo gli approcci descritti sopra

Si valuta imprecisione della stima della NMA come sopra

Se le due stime contribuiscono in egual misura

devo vedere se sono coerenti

se sono coerenti:

1.32 (0.98; 1.77)

Indirect: 1.35 (0.64; 2.86)

Network: 1.30 (1.08; 1.58)



valuto certezza di entrambe per tutte le dimensioni tranne imprecisione considero quella con certezza più alta

Si valuta imprecisione della stima della NMA come sopra

Se non sono coerenti

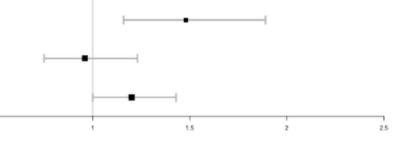
Procedo come sopra ma abbasso ulteriormente per incoherence

Direct:

1.48 (1.16; 1.89)

Indirect: 0.96 (0.75; 1.23)

Network: 1.20 (1.00; 1.43)



NMA-SoF table example 2

Estimates of effects, credible intervals, and certainty of the evidence for chemoprevention of colorectal cancer in individuals with previous colorectal neoplasia

BENEFITS

Patient or population: Individuals with previous colorectal neoplasia

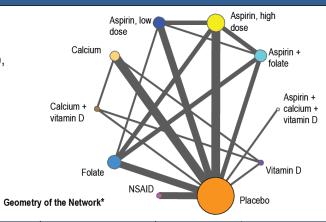
Interventions: Low and high dose aspirin, nonaspirin non-steroidal anti-inflammatory drugs (NSAIDs),

calcium, vitamin D, folic acid

Comparator (reference): Placebo

Outcome: Prevention of advanced neoplasia; range of follow up between three to five years

Setting: Outpatient



Bayesian NMA-SoF table

	al studies: 21 RCT	Relative effect** (95% Crl)	Anticipated absolute effect*** (95% Crl)			Certainty of	Ranking****	Interpretation
Tota	al Participants: 12088		Without intervention	With intervention	Difference	evidence	(95% Crl)	of Findings
•	Aspirin + calcium + vitamin D (1 RCT; 427 participants)	0.71 (0.18 to 2.49) Network estimate	74 per 1000¹	53 per 1000	21 fewer per 1000 (61 fewer to 110 more)	⊕⊕○○ Low Due to Imprecision ^{2, 5}	3 (1 to 10)	Probably inferior
•	Calcium + vitamin D (1 RCT; 1028 participants)	0.91 (0.52 to 1.63) Network estimate	74 per 1000¹	67 per 1000	7 fewer per 1000 (36 fewer to 47 more)	⊕⊕○○ Low Due to Imprecision ^{2,5}	6 (1 to 10)	Probably inferior
•	Aspirin + folate (2 RCT; 916 participants)	0.73 (0.43 to 1.19) Network estimate	74 per 1000¹	54 per 1000	20 fewer per 1000 (42 fewer to 14 more)	⊕⊕○○ Low Due to Imprecision ^{2,5}	4 (2 to 8)	Probably inferior
•	Aspirin, high dose (3 RCT; 917 participants)	0.81 (0.50 to 1.28) Network estimate	74 per 1000¹	60 per 1000	14 fewer per 1000 (37 fewer to 21 more)	⊕⊕○○ Low Due to Imprecision ^{2, 5}	5 (2 to 9)	Probably inferior

NMA-SoF table example 2

•	Aspirin, low dose (3 RCT; 823 participants)	0.71 (0.41 to 1.23) Network estimate	74 per 1000¹	53 per 1000	21 fewer per 1000 (44 fewer to 17 more)	⊕⊕○○ Low Due to Imprecision ^{2, 5}	3 (2 to 9)	Probably inferior
•	Nonaspirin NSAIDs (4 RCT; 3486 participants)	0.37 (0.24 to 0.53) Network estimate	74 per 1000¹	27 per 1000	47 fewer per 1000 (56 fewer to 35 fewer)	⊕⊕⊕ High.⁵	1 (1 to 2)	Definitely superior
•	Vitamin D (1 RCT; 764 participants)	1.19 (0.65 to 2.15) Network estimate	74 per 1000¹	88 per 1000	14 more per 1000 (26 fewer to 85 more)	⊕⊕○○ Low Due to Imprecision ^{3, 5}	9 (3 to 10)	Probably inferior
•	Calcium (3 RCT; 2503 participants)	1.00 (0.66 to1.52) Network estimate	74 per 1000¹	74 per 1000	0 fewer per 1000 (25 fewer to 38 more)	⊕⊕○○ Low Due to Imprecision ^{4, 5}	7 (3 to 10)	Probably inferior
•	Folate (3 RCT; 1224 participants)	1.32 (0.85 to 2.00) Network estimate	74 per 1000¹	51 per 1000	23 more per 1000 (11 fewer to 74 more)	⊕⊕○○ Low Due to Imprecision ^{2, 5}	9 (5 to 10)	Probably inferior
•	Placebo	Reference comparator	No estimable	No estimable	No estimable	Reference comparator	7 (4 to 9)	Reference comparator

NMA-SoF table definitions

GRADE Working Group grades of evidence (or certainty in the evidence)

High quality: We are very confident that the true effect lies close to that of the estimate of the effect

Moderate quality: We are moderately confident in the effect estimate: The true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different

Low quality: Our confidence in the effect estimate is limited: The true effect may be substantially different from the estimate of the effect

Very low quality: We have very little confidence in the effect estimate: The true effect is likely to be substantially different from the estimate of effect

Explanatory Footnotes

- ¹ Baseline risks (assumed control risk) obtained from the National Cancer Institute pooling project
- ²Very serious imprecision since 95% Crl crosses unity, and with wide credible intervals suggesting high possibility of harm.
- ³ Very serious imprecision since RR>1 (suggesting greater likelihood of harm than benefit), and with wide credible intervals).
- ⁴Very serious imprecision since RR is one (suggesting no evidence of benefit) and wide credible intervals suggesting high possibility of harm.
- Conceptually, there was no significant intransitivity, with comparable distribution of plausible effect modifiers across trials of different chemopreventive agents.

^{*} Lines represent direct comparisons

^{**} Estimates are reported as odds ratio. Crl: credible interval. Results are expressed in credible intervals as opposed to the confidence intervals (CI) since a Bayesian analysis has been conducted.

^{***} Anticipated absolute effect. Anticipated absolute effect compares two risks by calculating the difference between the risks of the intervention group with the risk of the control group.

^{****} Surface under the cumulative (SUCRA) ranking and credible intervals for efficacy are presented. Rank statistics is defined as the probabilities that a treatment out of n treatments in a network meta-analysis is the best, the second, the third and so on until the least effective treatment.

Buon lavoro e.... Tanti auguri ©